

Dynamics of Deterrence: A Macroeconomic Perspective on Punitive Justice Policy*

Bulent Guler
Indiana University

Amanda Michaud
Federal Reserve Bank of Minneapolis

October 22, 2024

Abstract

We argue that transitional dynamics play a critical role in evaluating the effects of punitive incarceration reform on crime, inequality, and labor markets. Individuals' past choices regarding crime and employment under previous policies have persistent consequences that limit their responsiveness to policy changes. We provide novel cohort evidence supporting this mechanism. A quantitative model of this theory, calibrated using restricted administrative data, predicts nuanced dynamics of crime and incarceration that are distinct across property and violent crime and similar to the U.S. experience after 1980. Increased inequality and declining employment accompany these changes, with unequal impacts across generations.

**Note:* Email: amanda.michaud@mpls.frb.org. For help with NACJD data, we thank the staff at ICPSR especially Arun Mathur, Brent Phillips, and Daric Thorne. A special acknowledgement to Paul Klein, who provided constructive criticism coupled with incessant encouragement that we complete this paper. For comments, we thank Jonathan Eaton, Giovanni Gallipoli, Erik Hurst, Karen Kopecky, Tatyana Koreshkova, Ricardo Lagos, Rasmus Lentz, Lance Lochner, Luigi Pistaferri, Ned Prescott, Victor Rios-Rull, Guillaume Rocheteau, Peter Rupert, Todd Schoellman, Pedro Silos, and Mark Wright; as well as seminar participants at Concordia University, Indiana University, and the Federal Reserve Banks of Atlanta, Chicago, Cleveland, Kansas City, & St. Louis; and participants at SED 2014, Midwest Macro, and LAEF Real Business CYCLE conference, and Notre-Dame Paella Workshop. The views expressed in this paper are those of the authors do not necessarily reflect the views of the Federal Reserve Bank of Minneapolis or the Federal Reserve System.

1 Introduction

Before the 1980s, the incarceration rate in the United States remained stable and comparable to other nations. Subsequently, the paths diverged.¹ A four-fold increase in the imprisonment rate from 1980 to 2000 made incarceration a common experience for less-educated men in the United States, despite recent modest declines.² It is widely accepted that from the 1980s onward, increased use of punitive incarceration for those arrested — not changes in crime or arrest rates — contributed to this divergence.³ There is little agreement, however, on the broad impacts of this substantial change in the justice system.⁴ Open questions range from assessing the effectiveness of these reforms in reducing crime to understanding the implications for economic outcomes and inequality.

We argue that understanding the dynamic consequences of policy reform — the changes slowly unfolding in the transitional decades following a policy change — is crucial for evaluating punitive incarceration policy.⁵ Criminal behavior is persistent at the individual level, on average.⁶ This leads to a weak deterrent effect of increased punitive incarceration in the short run, as the lingering consequences of past choices are difficult to reverse even when punishment becomes more severe. A temporary spike in incarceration can then occur amidst inelastic short-run behavior. If an incarceration experience increases future deviance through worse labor market prospects or the accumulation of criminal capital, this spike may translate into increased crime in the short run. As new cohorts born under the stricter policy reach their peak crime years, the full deterrent effect is finally realized, and both crime and incarceration fall in tandem. This pattern remarkably resembles the U.S. experience after 1980, particularly for property crime: a monotonic decrease in crime, alongside a rise and fall in incarceration (Figure 1 shows these patterns for each property and violent crime). In this paper, we study how these dynamic paths can stem from a single increase in punitive

¹See Burnham and Burnham (1999) for cross-country data and Hindelang (2016) for historical U.S. data.

²On a given day in 2008, an estimated 12.0% (37.2%) of white (black) males between the ages of 20 and 34 without a high school degree were incarcerated, (Pettit (2012)).

³Neal and Rick (2014) make this argument using the same administrative data as this paper. See also Blumstein and Beck (1999), Pfaff (2012), and Raphael and Stoll (2009) for theories of the underlying drivers, ranging from policy changes to career incentives of district attorneys.

⁴Bushway (2011) points out that in addition, little is known about which specific policies have been most influential.

⁵Analysis of the dynamic effects of policy changes given the dynamic nature of individuals' choices to participate in crime, appears little explored in the literature (McCrary (2010) provides a review). The closest related paper, İmrohoroğlu et al. (2004), compares property crime in the early 1980's with that in the late 1990's assuming full transition to a new steady state after policy change. A large literature estimates dynamic models of criminal behavior, but does not include policy changes.

⁶As many of half of the individuals released from prison in the U.S. will be reincarcerated within three years (calculated from the Department of Justice: Recidivism of Prisoners Released in 1994 data series).

policy.⁷

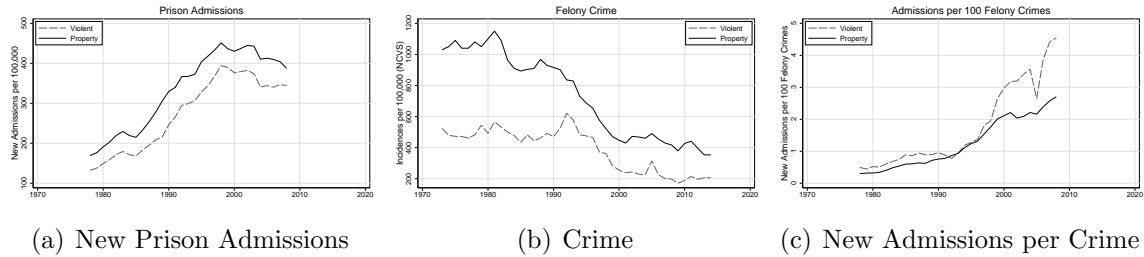


Figure 1: Trends in incarceration and crime. Authors’ calculations from NCRP, BJS, and NVS data. See appendix and Section 4 for details.

The unique cohort predictions of this theory are validated with a novel empirical strategy to separately identify age, cohort, and time effects. We find evidence of a “lost cohort” of individuals born in the mid-to-late 1960s – individuals in their 20s, the prime crime age, in the 1980s, have higher rates of prison admission and arrests *throughout their lives*, compared with those of the generations before them *and generations following*. This provides a critique of the criminal justice literature that attributes the increased average age of criminals to a fundamental shift in the age profile.⁸ We show that this is only partially a shift in the age profile, as also predicted by the theory we develop; it is also due to a cohort effect. The prevalence of cohort effects also highlights the importance of considering dynamics: the implication that the costs and benefits of blunt reforms are borne unequally across generations.

To investigate the dynamic consequences of punitive incarceration policy reform, we develop an overlapping generations model with several channels contributing to criminal persistence. Building on Becker (1968) theory of rational crime, where agents trade off labor market opportunities and criminal activities, we enrich the model with additional elements necessary to replicate the joint persistence of criminal behavior and labor market outcomes observed in data. First, human capital accumulates during employment and depreciates during non-employment. Second, criminal capital grows with engagement in crime and decays with age. Third, a criminal record, observable to employers, can limit employment opportunities. These ingredients lead to divergent paths of individuals’ employment and criminal propensities consistent with micro-data: widespread crime among the young, high

⁷This is a particularly important point, given the inference on the relationship between aggregate crime and incarceration featured in policy discourse. For example, Eisen and Cullen (2016) point out that “Imprisonment and crime are not consistently negatively correlated... This contradicts the commonly held notion that prisons always keep down crime.” We provide a model explicitly showing the flaw in applying causal interpretation to aggregate series in this way that goes beyond convoluting orthogonal factors.

⁸See Steffensmeier et al. (1989), Gottfredson and Hirschi (1990), and others.

recidivism rates, and limited crime-employment or crime-wage differentials.

We calibrate the model to quantitatively discipline the channels of criminal persistence by requiring it to match both cross-sectional and aggregate data. Our empirical strategy leverages an array of high-quality restricted administrative data from various sources, including the Survey of Inmates of State Correctional Facilities; a three year panel of parole officer data on over 12,000 individuals (Recidivism of Felons on Probation, 1986-1989); and the large-scale panel of annual prison censuses (National Corrections Reporting Program Data). This approach is distinct from prior micro-econometric and structural estimations that have typically relied on survey data from current and former inmates self-reporting their employment and criminal activity, which suffer from non-response, misreporting, and small sample sizes. In contrast, we utilize substantially larger and more reliable administrative data.⁹

Our main quantitative exercise evaluates the contribution of increased punitive incarceration to the U.S. prison boom and related outcomes. Our primary analysis considers property crime, which we later contrast with violent crime.¹⁰ For property crime, we simulate an increase in the probability of incarceration conditional on committing a crime from 0.3% to 2.1%, as estimated from U.S. data. We incorporate observed changes in real wages and estimated changes in returns to crime. The incarceration rate increases from 0.59% to 1.28% over the first 15 years, then declines over the next 30 years towards a new steady-state incarceration rate of 0.96%. Crime falls continuously over 30 years, from 1.35% to 0.31%, because of the immediate incapacitation of the most active criminals and more gradual deterrence effects on new generations' crime entry decisions. Furthermore, as is consistent with the data, crime becomes more concentrated among persistent career criminals. The model parsimoniously replicates, without targeting, the shape and magnitude of the non-monotonic cohort effects that motivated this study. It predicts that policy changes had permanent effects on inequality, with the employment gap for those with records steadily widening to a 2.5 percentage point deficit as fewer employers offer jobs to individuals with criminal records.

To complement the main exercise, we add several illustrative experiments and decompositions. A regression analysis on model data shows that aging is the largest factor deterring crime. Employment status has no significant impact, but a 10% increase in wage prospects lowers crime by 1.2%. Having a past prison experience is an important predictor of future

⁹The National Longitudinal Survey of Youth includes a panel of interviews of a two cohorts of individuals before, during, and after imprisonment. The sample reporting incarceration features fewer than 200 people, and these individuals have many non-responses.

¹⁰How do property and violent crime translate into understanding larger trends? More than 50% of prisoners have a conviction for violent crime. Only 16% of state prisoners are on drug charges, and 5-6% are nonviolent drug offenders. Sevigny and Caulkins (2004)

crime. Considering these factors in the transition, we find that harsher punitive policies lowered crime by counteracting trends of increasing criminal rewards and declining real wages. Finally, we study how the impact of punitive policy depends on the initial steady state. The marginal reduction in crime diminishes sharply when starting from more punitive initial policies. This is because in more punitive regimes, a larger share of crime reduction comes from incapacitation rather than deterrence, which affects dynamics since incapacitation is instantaneous, while deterrence can be delayed. Thus, the short-run elasticity of crime to policy approaches the long-run elasticity when starting from a more punitive initial stance. Repeating our analysis for violent crime provides an interesting comparison. The model estimation produces a calibration where violent crime resembles “crimes of passion”—it is more random and less persistent. Consequently, crime entry is less elastic to policy, and so dynamics become less important for the transition after a policy change.

These findings are not only important for accurately evaluating justice policies in real-time, but hold promise for improving their design. When crime is more persistent, as with property crime, crime reduction immediately after a policy change comes almost entirely from incapacitation effects, while deterrence effects on crime entry build over time. These conclusions should encourage the study of dynamic punitive policies specifying paths for multiple levers: the probability and duration of incarceration, as well as differential penalties for new and repeat offenders.

The literature on crime features few structural equilibrium approaches. Engelhardt et al. (2008) consider how the ability of employers to write efficient contracts tempers the labor market response to crime and vice versa. Huang et al. (2004) and Burdett et al. (2003) study interactions with the labor market in search frameworks. The most related papers are İmrohoroğlu et al. (2004), Fella and Gallipoli (2014) and Engelhardt (2010). İmrohoroğlu et al. (2004) quantify the contributions of changes in apprehension probability, labor markets, and population aging to the decline in property crime.¹¹ Fella and Gallipoli (2014) also consider property crime, but evaluate the impact of educational policy as well as punitive policy on crime. Engelhardt (2010) develops a model with rich heterogeneity to match the cross-sectional distribution of who commits property crime. Our work differs because we consider transitional dynamics.¹² Nonetheless, there are many similarities between our model and the ones in these papers: pecuniary considerations that differ according to life-cycle human capital growth and on employment status, and criminal capital or fixed heterogeneity

¹¹Similarly, Caucutt et al. (2021) study the effect of the War on Crime on the marriage gap between black and white men.

¹²Fu and Wolpin (2018) and Lochner (2004) are other prominent examples of structural models of crime that focus on long-run effects of policies on crime. Fu and Wolpin (2018) study the effects of policing on crime. Lochner (2004) studies the effects of education policies on crime.

to account for patterns of crime that pecuniary features alone cannot match within their respective frameworks. As will be clear, we place extra care in parsing those components of heterogeneity, as this is important for transition dynamics.

2 Importance of Criminal Persistence: Evidence from Cohorts

Our theoretical model studies how criminal persistence shapes dynamic responses to changes in punitive policy. In this section, we argue that criminal persistence generates two testable predictions related to cohort differences and changes in the age profile of crime. We then provide novel empirical evidence supporting these predictions.¹³

Criminal persistence relates to crime and incarceration dynamics through three potential channels. First, punitive policy – specifically, the probability of incarceration conditional on committing a crime – deters crime. Second, criminal behavior is persistent, but age eventually deters crime. Third, an incarceration experience affects the likelihood of future crime and incarceration.

The first and second channels cause cohorts who are at peak crime ages when punitive policy increases to have higher crime and incarceration rates throughout their lives than those of younger cohorts. The deterrence channel suggests that an increase in punitive policy causes people to choose lower criminal activity, all else equal. The persistence channel implies that individuals with no criminal history have a higher elasticity to policy changes than those with a criminal record. Newborns have zero criminal history and are most responsive to the policy change, whereas cohorts at peak crime ages are the least responsive.

Whether cohorts at peak crime ages when punitive policy increases differ from past cohorts depends on both the third channel (the criminogenic effect of incarceration) and the elasticity of criminal behavior to the policy change. The third channel implies that the peak cohort will have even higher crime and incarceration rates than past cohorts because of their increased likelihood of incarceration, subsequent criminal engagement upon release, reincarceration, and so on. The criminogenic effect of prison also raises crime in later years, providing an additional prediction: age curves should become flatter, with higher criminal activity at older ages relative to younger ages. However, these effects can be offset if crime choices are highly inelastic. In the limit, if nobody’s criminal behavior changes, all cohorts would look the same.

¹³We will use our full structural model to isolate these effects as well and compare them with the direct empirical estimates.

Testing these predictions requires disentangling age, time, and cohort effects in a non-colinear (additive) manner.¹⁴ Our assumptions are as follows: the cohort effect is a level effect that shifts the entire age profile up or down, capturing cohort-specific factors that affect criminal behavior throughout cohort members’ lives. The age effect is a growth/decay rate that defines the shape of criminal behavior over the life course, regardless of level.¹⁵ The time effect enters in two ways. The first is a linear component that increases/decreases crime or incarceration for all cohorts alive, scaled by their age effect. The second is a component that multiplies the age effect, allowing the decay of crime over the life course to vary over time, which would be the case if prisons are criminogenic. This second component tests the prediction of flattening age profiles.

Formally, the non-linear least squares model is as follows:¹⁶

$$I_{a,c,t} = (\beta^T \mathbf{D}^T + \beta^C \mathbf{D}^C) * (\beta^A \mathbf{D}^A * \beta^Y \mathbf{D}^T) + \epsilon_{a,c,t}$$

$$\text{st } \beta^Y = 0 \quad \text{if } a < 26$$

The dependent variable is the prison admission rate of cohort c , which is aged a at time t . The independent variables D^T , D^C , and D^A are respectively dummies for time, age, and cohorts.¹⁷ We assume β^Y is zero before the peak of the life-cycle incarceration curve, so it captures only the flattening of the life-cycle profile and how it changes over-time.¹⁸

Figure 2 presents the estimated cohort and time coefficients for prison admissions on primarily property (top) and violent (bottom) crime charges. Cohort effects are distinct, significant, and larger than time effects in the case of property crimes. The cohort effect is largest for those born in the early-to-mid 1960s. This is consistent with the view that punitive incarceration policy increased most sharply in the 1980s, when these cohorts would have been in their peak criminal careers. Interestingly, the model infers that time effects, at least in levels, did not play a significant role in the rise in incarceration related to property

¹⁴This is a contribution to the criminal justice literature, which has focused mostly on the changing age-structure of prison admissions, something we demonstrate can be attributed partially to cohort effects.

¹⁵This strategy relates to Schulhofer-Wohl and Yang (2016) and (Lagakos et al., 2016). We overcome co-linearity by placing more structure on the nature of the age effects. We also directly address the issue raised in (Schulhofer-Wohl and Yang, 2016) advocating that the age effect may be changing over time and cohorts because time effects have both a level and life-cycle growth rate component.

¹⁶Here, we will present results for incarceration for property and violent crime— specifically prison admissions. The Online Appendix contains results for crime rates as well as prison admissions for other types of crime

¹⁷We measure cohorts and time in five year intervals.

¹⁸This relates to Lagakos et al. (2016) who advocate using theory to identify where the age effect is negligible. Here, we assume that age effects are negligible in early life, compared with time and cohort influences. We provide estimations in the Online Appendix for the case where $\beta^Y = 0$ for all ages, which means the age effect is not allowed to change over time. This unsurprisingly fits the data worse but also works in favor of our hypothesis by increasing the magnitude of the cohort effects.

crimes. Cohort effects are also significant for violent crime but are not humped-shaped, despite the increase in punitive policy. They are also of similar size to the time effects. The simple model would interpret these facts as suggesting that (1) punitive incarceration policy deterred both property and violent crimes but deterred property crime more; (2) criminal persistence caused the deterrent effect to materialize slowly as new cohorts are born for both types of crime; and (3) prison is criminogenic for property crime but not violent.

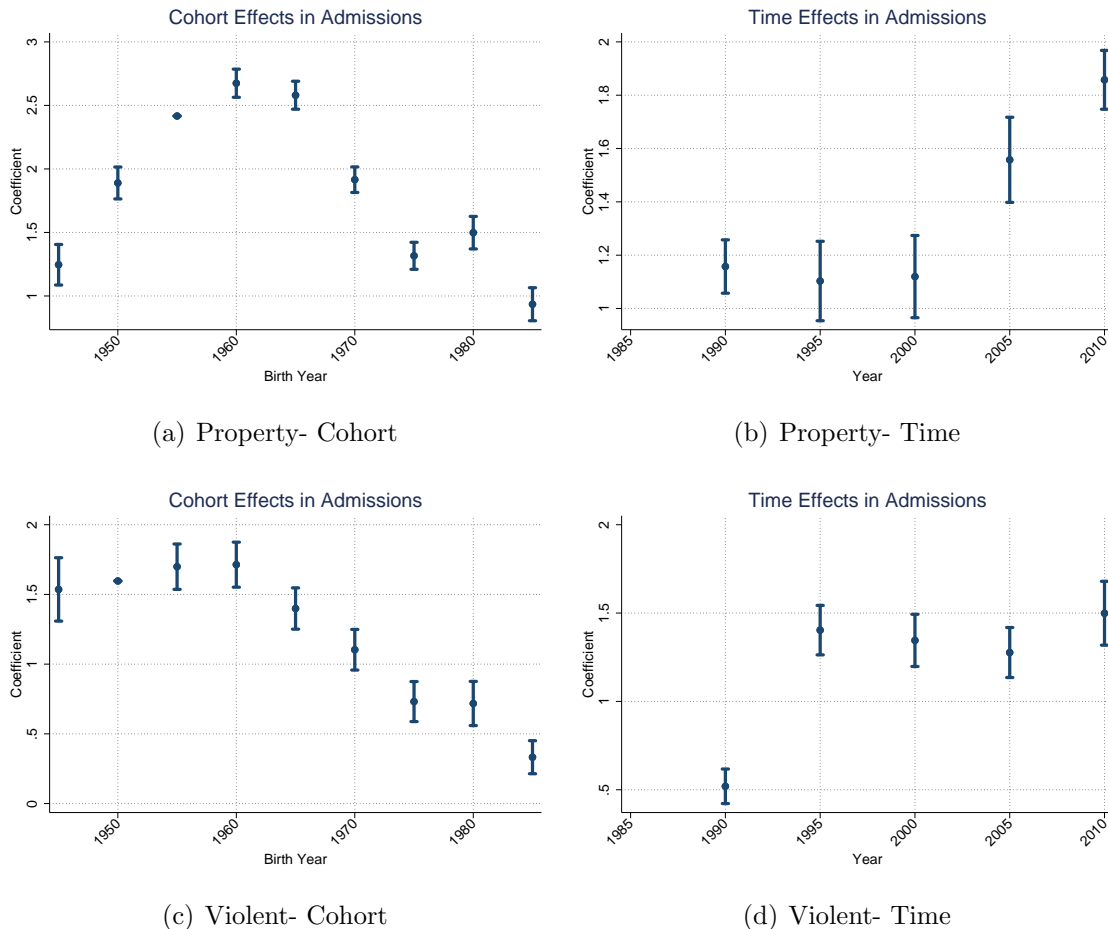


Figure 2: Estimated Cohort and Time Effects. Top: property crime prison admissions. Bottom: violent crime prison admissions.

While our theory is not the only one that can generate cohort effects, other explanations for why the 1960s cohorts have uniquely high crime and incarceration rates – such as the lead hypothesis, trailing the baby boomers, or being the right age to enter the crack trade – point to temporary impacts affecting a single cohort. In contrast, the idea that these cohort effects are driven, at least in part, by a permanent shift in policy uniquely implies permanent changes to the age profile when criminal behavior is persistent. It predicts that even later-born cohorts with less overall crime in their lifetime should see increased crime

and incarceration at older ages, relative to when they are young. In other words, our theory generates unique predictions for the intensive and extensive margins of crime.

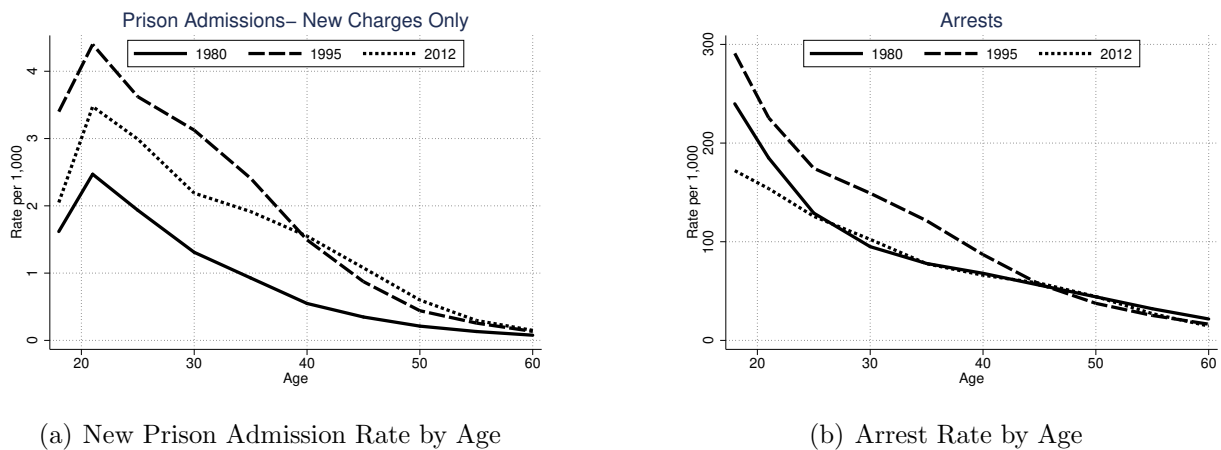


Figure 3: **Permanent and Cohort Shifts in Age Profiles (Data):** Prison admissions from National Corrections Reporting Program Data. Arrests from FBI crime reports accessed through the Bureau of Justice Statistics

3 Quantitative Model

We present a quantitative model built on Burdett et al. (2003) and Engelhardt et al. (2008) to study how punitive incarceration policy affects crime rates, incarceration rates, and equilibrium labor market outcomes.

Time is continuous. The economy is populated by a continuum of finitely-lived ex-ante identical individuals and identical firms. Individuals have linear preferences over consumption and discount the future at rate $0 < r < 1$. At any point in time, individuals experience one of three labor market statuses: (i) employment, (ii) unemployment, or (iii) incarceration.

3.1 An Individual’s Problem:

An individual is characterized by five state variables: age, employment status, human capital, past incarceration records, and criminal capital. The first three sources (age, employment, human capital, and records) provide observable links between the model and salient cross-sectional variation in criminality in the data. In the data, crime is more concentrated in people who are younger, are unemployed, have lower earnings, and have criminal records,

and so the economic mechanisms in the model should capture these dimensions. The final source, criminal capital, is an unobserved residual used to generate observed criminal persistence within individuals that cannot be provided by the first three ingredients. Together, these ingredients allow the model to match both the extensive (cross-section) and intensive (individual persistence) margins of crime both, of which contribute critically to the dynamics of deterrence.

Age takes a finite number of values: $m \in M = \{1, \dots, \bar{m}\}$. Individuals become age $m + 1$ at the poisson rate ϑ^m .¹⁹ When individuals at the maximum age, \bar{m} , receive an aging shock, they exit the economy, receive zero continuation utility, and are replaced with age 1 individuals who start life with the lowest skill level and are initially unemployed.

Employment opportunities arrive at the poisson rate λ_w . All jobs are identical. Upon receiving a job opportunity, the unemployed individual can either accept the offer or reject it. If they accept, they become employed and receive a flow wage proportional to their human capital (productivity) level: hw , where w is the piece rate and h is their current human capital. Employed individuals receive a job separation shock at poisson rate δ , at which point they become unemployed. Unemployed individuals receive flow consumption bwh .

Each individual is endowed with an identical initial human capital level. Human capital changes at the poisson rate ψ and evolves according to labor status dependent function $f_j(h)$ given current human capital level h . That is, $h' = f_j(h)$, where $j \in \{e, u, p\}$.

Rational crime opportunities arrive at age specific rates η^m . Each opportunity that arrives presents an iid draw of an instantaneous reward κ , separable in utility from consumption and drawn from a common distribution H . Upon receiving a crime opportunity, the individual sees the particular κ drawn and decides whether to commit the crime. Individuals who commit crimes are caught with probability π .

Criminal capital takes two values: low (lc) and high (hc). High criminal capital individuals receive additional crime opportunities. These opportunities arrive at rate η_a^{hc} and must be committed without any additional benefit to the individual. All individuals are born with low criminal capital. Low criminal capital changes to high with probability ν when a crime is committed. High criminal capital depreciates to low with age-specific probability ζ^m .

Individuals who commit crimes are caught and incarcerated with probability π . Incarcerated individuals receive zero flow utility. They are released from prison to unemployment at rate τ . Individuals who have been to prison are distinguished to employers by a criminal record. We allow separate job markets: one for individuals who have never been incarcerated, called *non-flagged* individuals, and another for individuals who have been incarcerated at

¹⁹Stochastic aging is a standard method of reducing the state space (in this case to 3 age groups instead of 2392 age-weeks) to make the computation feasible. It is not a source of meaningful economic risk.

least once, called *flagged* individuals. We denote k as the flag type, and $k = 0$ refers to non-flagged, whereas $k = 1$ refers to flagged individual. This feature is included to capture the market segmentation that arose both from occupational restrictions for ex-felons and from employer use of criminal records in screening during our study period.²⁰ In keeping with realism, employers cannot observe certain individual characteristics, like criminal capital. They can, however, use the criminal record flag to statistically deduce criminal propensity.

We denote V_p , V_u , and V_e as the value of an incarcerated, unemployed, and employed individual, respectively. The recursive formulation of an incarcerated individual's problem is

$$rV_p(h, i, m) = \psi \int (V_p(h', i, m) - V_p(h, i, m)) f_p(h') dh' + \zeta^m (V_p(h, lc, m) - V_p(h, i, m)) \\ + \tau (V_u(h, i, 1, m) - V_p(h, i, m)) + \vartheta^m (V_p(h, i, m + 1) - V_p(h, i, m)) \quad (1)$$

where $i \in \{lc, hc\}$ is the criminal capital level, h is the current human capital level, and m is the current age of the individual. The first term on the right-hand side reflects the change in the value upon receiving human capital shock, the second term captures the change in value upon receiving rehabilitation shock, the third term captures the change in value upon receiving the prison exit shock, and the final term reflects the change in value upon receiving the age shock. We assume $V(\cdot) = 0$ when $m = \bar{m} + 1$. Implicit in this formulation is that as a normalization, the incarcerated individual receives 0 flow utility while incarcerated.

The problem of an unemployed individual is

$$rV_u(h, i, k, m) = bwh + \psi \int (V_u(h', i, k, m) - V_u(h, i, k, m)) f_u(h') dh' + \\ \lambda_w^{k, m} (V_e(h, i, k, m) - V_u(h, i, k, m)) + \vartheta^m (V_u(h, i, k, m + 1) - V_u(h, i, k, m)) + \\ \zeta^m (V_u(h, lc, k, m) - V_u(h, i, k, m)) + \eta_a^k (V_{up}(h, i, k, m; 0) - V_u(h, i, k, m)) + \\ \eta^m \int \max \{V_{up}(h, i, k, m; \kappa) - V_u(h, i, k, m), 0\} dH(\kappa) \quad (2)$$

where $i \in \{lc, hc\}$ is criminal capital and $V_{up}(h, i, k, m; \kappa) = \pi (V_p(h, 1, m) \nu + V_p(h, i, m) (1 - \nu)) + (1 - \pi) (V_u(h, 1, k, m) \nu + V_u(h, i, k, m) (1 - \nu)) + \kappa$ denotes the value upon committing a property crime with reward κ . It includes the probability of incarceration π and probability of gaining high criminal capital ν , each associated with committing the crime. With probability $(1 - \pi)$, the individual is not caught, but is still subject to change in criminal capital. The first term

²⁰Harmonized electronic records across jurisdictions began to be available in the mid-1990s, however, analyzing the impacts of record access is non-trivial because access remained highly variable across states for over a decade. Also, explicit records are unlikely to be the only avenue through which criminal history could be ascertained. These issues are beyond the scope of this paper.

is the flow benefit of unemployment. The rest of the terms capture the change in value upon the receiving human capital shock, an employment opportunity, an aging shock, a rehabilitation shock, an irrational crime opportunity (if they are high criminal capital type), and a rational crime opportunity, respectively.²¹

The recursive formulation of an employed individual's problem is

$$\begin{aligned}
rV_e(h, i, k, m) = & \quad wh + \psi \int (V_{eu}(h', i, k, m) - V_e(h, i, k, m)) f_e(h') dh' + \\
& \delta (V_u(h, i, k, m) - V_e(h, i, k, m)) + \vartheta^m (V_e(h, i, k, m+1) - V_e(h, i, k, m)) + \\
& \zeta^m (V_e(h, lc, k, m) - V_e(h, i, k, m)) + \eta_a^k (V_{ep}(h, i, k, m; 0) - V_e(h, i, k, m)) + \\
& \eta^m \int \max \{V_{ep}(h, i, k, m; \kappa) - V_e(h, i, k, m), 0\} dH(\kappa)
\end{aligned} \tag{3}$$

where $V_{ep}(h, i, k, m; \kappa) = \pi (V_p(h, 1, m) \nu + V_p(h, i, m) (1 - \nu)) + (1 - \pi) (V_e(h, 1, k, m) \nu + V_e(h, i, k, m) (1 - \nu)) + \kappa$. The first term is the flow wage income, which is proportional to the human capital. The rest of the terms capture the change in value upon receiving human capital shock, job separation shock, aging shock, rehabilitation shock, irrational crime opportunity, and rational crime opportunity, respectively.

The only decision rule of the individual is the crime decision, which we denote as I_u for the unemployed and I_e for the employed:

$$I_j(h, i, k, m; \kappa) = \begin{cases} 1 & \text{if } V_{jp}(h, i, k, m; \kappa) \geq V_j(h, i, k, m) \\ 0 & \text{o.w} \end{cases}$$

where $j \in \{u, e\}$.

3.2 Matching

Employers create jobs conditional on individuals' observable traits: their criminal record flag k and their age m .²² This segments the economy into $2M$ labor markets.²³ All labor markets are modeled as in Pissarides (1985). Employers with vacant jobs and unemployed workers meet randomly according to a matching function $M(u_{km}, v_{km})$, where u_{km} and v_{km} are the number of unemployed workers and vacant jobs for individuals with flag type k and age m . The matching function is strictly increasing in both terms and has constant returns

²¹Individuals with low criminal capital do not receive irrational crimes crime opportunities; that is, $\eta_a^{lc} = 0$.

²²Age has been shown to be an important screening mechanism when criminal records are not available (see, e.g., Doleac and Hansen (2020)).

²³By assuming workers only search within markets for their age/flag type, we are silent about issues of hold-up problems or commitment if a worker is matched with a job in a market that is different than their age/flag type.

to scale. The job arrival rate for workers can be expressed as:

$$\lambda_w^{k,m} = M(u_{km}, v_{km})/u_{km} = M(1, v_{km}/u_{km}) = M(1, \theta_{km}), \quad (4)$$

where θ_{km} is the market tightness for type- km jobs. Similarly, vacant job filling rate for firms can be expressed as

$$\lambda_f^{k,m} = M(u_{km}, v_{km})/v_{km} = M(u_{km}/v_{km}, 1) = M(1/\theta_{km}, 1) = \lambda_w^{k,m}/\theta_{km}. \quad (5)$$

3.3 A Firm's Problem:

Firms choose to post vacancies in each labor market so long as the net expected value is positive. The flow cost of posting a vacancy is c . The expected revenues from posting a vacancy are equal to the expected revenues from a match discounted by the equilibrium match arrival rate. A match with a worker with human capital level h produces $y = h$. The wage is assumed to be a constant fraction of the output of the match, and so the firm's flow profits equal $h(1 - w)$.²⁴ The match dissolves if either (i) the worker receives a separation shock; or (ii) if the worker commits a crime and gets imprisoned. Firms use rational expectations in line with the equilibrium distribution of human capital and criminal capital of workers searching in each particular age-cross-criminal record flag market to compute both the expected revenues of a match.²⁵

We denote the value of a filled job as J_f . The recursive formulation of a firm's problem is

$$\begin{aligned} rJ_f(h, i, k, m) = & h(1 - w) + \delta(V_f(k, m) - J_f(h, i, k, m)) + \psi_e \int (J_f^e(h', i, k, m) - J_f(h, i, k, m)) f_e(h') dh' + \\ & \vartheta^m (J_f(h, i, k, m + 1) - J_f(h, i, k, m)) + \eta_a^k (1 - \pi) (J_f^{pe}(h, i, k, m; 0) - J_f(h, i, k, m)) + \\ & \eta^m (1 - \pi) \int (J_f^{pe}(h, i, k, m; \kappa) - J_f(h, i, k, m)) dH(\kappa) + \zeta^m (J_f(h, lc, k, m) - J_f(h, i, k, m)) \end{aligned}$$

where J_f^{pe} is defined as

$$J_f^{pe}(h, i, k, m; \kappa) = \begin{cases} J_f(h, 1, k, m)\nu + J_f(h, i, k, m)(1 - \nu) & \text{if } V_e(h, i, k, m; \kappa) \geq V_e(h, i, k, m) \\ V_f(k, m) & \text{o.w.} \end{cases}$$

²⁴Nash bargaining is an alternative wage protocol but bargained wages create an odd outcome in models of rational crime: more criminally active individuals have better outside options and bargain higher wages. This tends to lead to equilibrium outcome that contradicts salient features of the data.

²⁵Any match with negative expected revenues is not formed, and the value equals zero.

The value of a vacant job is defined as

$$rV_f(k, m) = -c + \lambda_f^{k,m} \int (J_f(h, i, k, m) - V_f) d\Gamma_u(h, i|k, m) \quad (6)$$

where Γ_u is the marginal cumulative density function of unemployed over human capital and criminal capital conditional on observable prison flag, k , and age, m .

3.4 Definition of a Stationary Competitive Equilibrium:

A competitive stationary equilibrium is a set of value functions V_p , V_u , V_e , J_f and V_f ; individuals' crime policy functions I_u and I_e ; market tightness for each submarket θ_{km} ; job arrival rate for workers in each submarket λ_w^{km} ; worker arrival rate for firms in each submarket $\lambda + f^{km}$; and a stationary distribution of individuals Γ such that the following conditions hold:

1. Policy functions I_u and I_e solve the individual's problem characterized in equations 1-3 taking job arrival rates λ_w^{km} as given. Value functions V_p , V_u and V_e are the associated value functions to these problems.
2. The firm's value functions J_f and V_f solve equations 6 and 6 taking worker arrival rates λ_w^{km} for each k and m , individual decision rules I_u and I_e , and the stationary distribution of individuals μ as given.
3. There is free entry: $V_f(k, m) = 0$ for each k and m .
4. The distribution is stationary and consistent with individuals' decision rules:

$$\Gamma = T(\Gamma)$$

where T is an operator mapping the current distribution to the future distribution given individuals' decision rules and law of motion for exogenous variables.

Individuals' crime policy functions take a reservation form. Conditional on one's state, there is a unique crime reward above which all crimes are committed and below which no crimes are committed.²⁶

²⁶This follows from strict monotonicity of the value functions. These proofs are standard and omitted (Rogerson et al. (2005)).

Mechanisms: Policies and Outcomes in a Stationary Equilibrium. The impact of a more punitive criminal policy, an increase in the probability of imprisonment for a crime, can be understood through several effects. The first is the deterrent effect through individuals' choices. As the probability of incarceration increases, all individuals, regardless of their status (summarizing their history), choose to commit less crime. The second is the arithmetic effect of an increase in the probability of imprisonment. If the increase in probability of getting imprisoned dominates the deterrent effect, then the incarceration rate increases. A third effect depends on how firms respond in equilibrium. If an increase in incarceration probability decreases the expected profits to a firm from hiring a worker, firms respond by posting fewer vacancies which results in lower job arrival rates for individuals. This equilibrium effect counters the deterrent effect of policy by increasing inducing both unemployed and employed workers to choose more crime.

4 Calibration and Estimation

We calibrate our model so that the initial steady state replicates empirical moments from the late 1970s and early 1980s. This choice is motivated by the prior century of comparably stable rates.²⁷ Some parameters are directly calibrated but most are jointly estimated to minimize the distance between the model and data statistics. The population of interest in the data is men with a high school degree or less. This text focuses on the calibration of the model for property crimes only. We produce a calibration for violent crimes in the Online Appendix.

4.1 Sources of Criminal Justice Data.

We calibrate our model to replicate moments from several sources of criminal justice data. This approach leverages varied, large, and representative administrative datasets, contrasting with prior studies often relying on small, self-reported samples from the NLSY or data from a single state or local agency.

Consistent nationwide data on prison admissions and criminal records are a challenge. Historically, annual records on prison admissions at the institutional level and individual criminal histories were collected by subgovernmental units responsible for particular justice system components, and each followed their individual conventions. Some improvement followed the 1993 Brady Act, which mandated background checks for some firearms purchases,

²⁷See the Online Appendix for a plot. Indeed, rates were so remarkably stable across space and time that a theory of a “natural rate” of incarceration was prominent for many decades, (Blumstein and Cohen (1973).

but overall, nationally aggregated data are collected by subnational authorities and should be viewed critically for irregularities.

The Bureau of Justice Statistics (BJS) estimates national prison admissions, stocks, and releases using data from the National Corrections Reporting Program (NCRP). The NCRP is a restricted-access dataset of offender-level data submitted to the BJS by state justice departments. As (Neal and Rick, 2014) note, the data require careful vetting. We clean the data by first dropping states in which the inflows, outflows, and stocks are not internally consistent, following (Neal and Rick, 2014). However, our interest in distinguishing property, violent, and other offenses necessitates additional consistency checks at the offense category level. This leaves us with data on 12 states, accounting for 42%-60% of all prison admissions over our period of interest, which exhibit trends similar to national BJS estimates (as shown in the Online Appendix). Finally, we perform additional checks for reliability, including investigating large growth or decline in admissions and more.²⁸ Finally, we perform additional reliability checks, interpolating outlier years instead of dropping entire states. These data consistently report offenders' age and gender, allowing us to restrict our sample to males and compute age group statistics where appropriate.²⁹

The Recidivism of Prisoners Released Series provides data on prisoner outcomes in the three years following release. These restricted offender-level administrative data include a representative sample of 16,000-38,624 prisoners released from states with large prison populations in the survey year, conducted every 11 years. We use the 1983 series to compute baseline recidivism statistics. Later surveys validate the model's predictions, but restricted 2005 micro-data are unavailable. Instead, we secured restricted micro-data from "Criminal Recidivism in a Large Cohort of Offenders Released from Prison in Florida, 2004-2008," which contains over 156,000 offenders. We verify these data align with the published 2005 BJS statistics, with a 3-year recidivism rate of 36% in the Florida data and 36.1% in the BJS data.

The Survey of Inmates of State Correctional Facilities provides data on labor market and personal characteristics at the time of offense for convicted prisoners. We use the 1979 survey of 12,000 inmates in 300 state institutions to calibrate the model. Nearly all state prisoners serve convicted sentences for one or more felonies. State prisoners also consistently comprise over 80% of all convicted prisoners, with the remainder mostly in federal institutions.

We follow a few general principles in categorizing crimes and convictions. When measuring crimes, we restrict our data to offenses likely charged as felonies, as individuals are rarely

²⁸Outliers aren't the only problem. We found a case where a state simply sent the same data to the BJS several years in a row!

²⁹Education is not reliably reported in these data and so we do not restrict our sample on the basis of education.

imprisoned for misdemeanors. To classify the offense for prison admission, we employ two strategies. When computing aggregate statistics such as total property crime admissions, we include all associated offenses. For example, if an individual’s main charge is violent but includes three property offenses, we count three property admissions. This approach accurately assigns the probability of incarceration to a single crime occurrence in victimization data.

This study focuses on property crime, with additional results for violent crime. Property and violent crimes are almost always victim-based, which facilitates estimation of a key policy parameter: the probability of incarceration per crime committed.³⁰

4.2 Externally Calibrated Parameters

The time period is set to be one week. Individuals go through three stages of life ($M = 3$): youth, middle age, and old age. On average, young individuals live for 7 years (between ages 18 and 24), middle-aged individuals live for 10 years (between ages 25 and 34), and old individuals live for 30 years (between ages 35 and 64).³¹ We set $r = 0.1\%$ to provide an annual discount factor of 0.95. We set the prison exit probability to 0.019, which implies 12 months of prison time on average, consistent with Both Raphael and Stoll (2009).³²

The matching function follows Shimer (2005)

$$M(u, v) = \chi u^\varphi v^{1-\varphi}$$

where u is the unemployment rate and v is the vacancy rate. As in Shimer (2005), we set the flow utility of unemployment b to equal 40%, the matching function curvature φ to 0.72, and the matching function constant χ to 0.14. We set the wage to be 50% of the productivity of the worker.³³

The incarceration probability upon committing a crime is set to $\pi = 0.003$. This value matches our calculation of new prison admits for property crime estimated from NCRP’s NPS restricted micro-data divided by number of property crimes estimated from the National Crime Victimization Survey (NCVS) for 1979-1980.³⁴

³⁰Details about crime counts and classification in the NCVS are in the Online Appendix.

³¹These average lifetimes for each age group imply the stochastic aging probabilities of $\vartheta_y = 0.00275$, $\vartheta_m = 0.00192$, and $\vartheta_o = 0.00064$ for the young, middle-aged, and old, respectively.

³²Raphael and Stoll (2009) and Neal and Rick (2014) show that the median prison time served has remained reasonably constant over time, whereas the average duration has increased because of the extreme tail (life sentences, etc.).

³³This value is inconsequential. The more important assumption is that workers with a higher outside option do not bargain higher wages.

³⁴See the Online Appendix for an extended discussion on how alternative measures of crime affect the

Table 1 shows the externally calibrated parameter values of the model.

Preset Parameters		
Parameter	Explanation	Value
ϑ_y	aging prob - young	0.00275
ϑ_m	aging prob - middle	0.00192
ϑ_o	aging prob - old	0.00064
τ	prison exit prob	0.019
r	discount factor	0.001
b	unemployment benefit	40%
φ	matching function curvature	0.72
χ	matching function constant	0.14
w	wage share	0.5
π	arrest probability	0.003
ψ	human capital shock arrival rate	1/52

Table 1: Externally Calibrated Parameters

4.3 Internally Calibrated Parameters

The remaining parameters in the model are jointly calibrated by minimizing the percentage deviation of the model-generated moments from their analogous data moments.³⁵ We explain our choices of moments to match below.

Labor Market Parameters: The employment rate is determined in equilibrium, in part by the decisions of workers and firms. Two parameters are also important: the exogenous job separation rate and the vacancy cost.³⁶ Targets for these parameters are the average employment rate and unemployment duration of men between the ages of 18 and 34, without a high school degree in 1980-83. We choose this demographic because they have the highest crime rates in the data. The estimated vacancy cost equals about one year of the average annual income in the economy. The calibration chooses a high vacancy cost to match the unemployment rate of 23.2% since we assume all non-employed workers are searching in unemployment.

time series of π . The appendix addresses why we do not use Uniform Crime Reports (UCR).

³⁵Specifics on the objective function, weighting matrix, and computation algorithm of the estimation process can be found in the Online Appendix, along with graphical relationships between individual parameters and moments.

³⁶The exogenous job separation rate cannot be set directly, because some matches dissolve endogenously when a worker is admitted to prison for a crime.

Human Capital Parameters: The human capital shock is chosen to arrive at a Poisson rate $\psi = 1/52$, without loss of generality. Upon receiving the shock, the log of the human capital follows an AR(1) process:

$$\log h' = f_{i,m}(h) = (1 - \rho_h)\mu_h^{i,m} + \rho_h \log h + \epsilon_h$$

where ρ_h is the persistence of the process, ϵ_h is a Gaussian white noise with variance σ_h^2 , and μ_h^i is the unconditional mean of the log human capital conditional on employment status $i \in \{e, u, p\}$, which captures the potential scarring effects of unemployment and incarceration. Scarring effects of unemployment and incarceration are assumed to be age-independent, and the mean for the first age group is normalized: $\mu^{e,1} = 0$, which implies an average human capital of $h = 1$. The remaining parameters of the process are estimated using indirect inference. The auxiliary model is the following Mincer regression run on both our NLSY 1979 sample and on data simulated in our model:

$$\ln(w_{it}) = \alpha + \beta^M I(A_{it} = 2) + \beta^O I(A_{it} = 3) + \beta^N N_{it} + \gamma_i + \epsilon_{it} \quad (7)$$

For an individual i at time t , w_{it} is the observed wage; I is the indicator function; A is an age bin; N is the months of non-employment, including unemployment and non-participation in the past year; γ_i is an individual fixed effect; and ϵ_{it} is a residual.³⁷ Given the shock arrival rate, ρ_h and ϵ_h are chosen to replicate estimates of the annual persistence and standard deviation of the residuals in the NLSY sample estimation. These statistics are 0.96 (persistence) and 0.2 (standard deviation), which are within the range of standard estimates used for men with a high school degree or less in the literature.³⁸ The three remaining parameters are estimated to minimize the distance between the coefficients on the age and last year non-employment indicators in the model and in the data, where prison counts as non-employment. These parameters are: $\mu^{e,2}$, $\mu^{e,3}$, and $\mu^{u,1}$.

Crime Parameters: Data moments on incarceration and recidivism rates serve as calibration targets to inform arrival rates of crime opportunities. The crime arrival rate for young and middle-aged individuals is assumed to be the same because first-time incarceration rates of young and middle-aged individuals are similar to each other in the data. By contrast, the

³⁷Since there is no ex-ante heterogeneity among individuals, we omit the fixed effect in the regressions for the simulated data. The model is weekly, but we store the information to construct the panel data at monthly frequency, as in the NLSY. Further details of the NLSY implementation and results can be found in the Online Appendix.

³⁸Storesletten et al. (2004) report higher variance for men with lower levels of education, with a range of 0.16-0.2 in a collection of similar studies.

first-time incarceration rates for those over age 34 is near zero in the data (<1%, authors' calculations from NACJD data), and so the crime arrival rate for old individuals with low criminal capital is calibrated as 0.

The share of the population with high criminal capital is crucial in determining the extent of recidivism (intensive margin) versus the extent of crime in the cross-section (extensive margin) in the economy. In other words, is crime done mostly by a few individuals who commit crimes frequently or by many individuals who commit crimes infrequently? This distinction helps distinguish two parameters related to criminal capital process: the probability of gaining high criminal capital after committing a crime, ν , and the additional crime arrival rate for high criminal capital individuals, η_a^{hc} . Without additional crime opportunities, the criminality of the high criminal capital types would be close to that of the general population, whereas crime is more concentrated in a few individuals in the data. Statistics on recidivism to prison are informative about the share of high criminal capital types and the additional crimes they commit. We add to our estimation targets the one-year re-imprisonment rate *on new charges* for the released prisoners.³⁹ This rate is 19.9% in the 1983 BJS Recidivism of Prisoners Released Study (Bureau of Justice Statistics (2011)) for young and middle-aged individuals.

The fraction of prison admits with prior incarceration experience is a complementary target. In the 1979 National Prison Survey, 64.2% of property criminals had been to prison before.⁴⁰ In the model, the probability of gaining criminal capital, ν , is a crucial parameter to capture this fact. If $\nu = 0$, crime will be more widespread among the population, whereas as ν becomes larger, crime will be concentrated among a few individuals. Together, the probability of gaining high criminal capital and the higher arrival rate of crime for this type determines the size of the population with higher than average crime rates. This relationship is a key feature of the data.

Crime rewards are drawn from exponential distribution with mean μ . We use the ratio of average wage of criminals to the average wage in the economy to discipline these parameters. If the average reward gets higher, the incentives for committing a crime will be smaller among high and low wage individuals, and the wage ratio will be smaller. The data counterpart of this wage ratio is the earnings prisoners reported in the month of the crime they were incarcerated for relative to those in the NLSY79, a ratio of 0.86 using NPS 1979 data.

³⁹These rates are calculated using the BJS Recidivism of Prisoners Released Series (Bureau of Justice Statistics (2011)). We take care to include only those re-imprisoned who are convicted of a new felony charge. This excludes those re-incarcerated in jails or re-imprisoned for violations of their parole, probation, or other conditions of release, in order to be consistent with the concept of incarceration and crime used in the model and in targets from other datasets.

⁴⁰Authors' calculation from the National Prison Survey 1979.

Parameter	Explanation	Value
η^1	crime arrival rate	0.07%
c	vacancy cost	80.17
δ	separation shock	1.39%
$\mu^{e,2}$	human capital mean-middle employed	0.11
$\mu^{e,3}$	human capital mean-old employed	0.16
$\mu^{u,1}$	human capital mean-nonemployed	0.14
ζ^3	rehabilitation shock	0.39%
ρ_h	human capital persistency	0.94
σ_h	human capital shock std	0.25
ν	prob of being high criminal	0.23
$\eta_a^{1,hc}$	high criminal crime arrival rate	1.57
μ^k	mean crime reward	0.61

Notes: The table shows the internally calibrated parameters of the model. See the main text for a discussion of the explanation of these parameters, and how they are identified in the model.

Table 2: Calibrated Parameters

Table 2 shows the calibrated parameters. Table 3 shows the performance of the model in matching the moments targeted. The model does a satisfactory job in capturing the moments targeted in the calibration.

5 Steady-State Analysis.

To understand the dynamics of deterrence following a change in punitive policy, we must first understand what determines crime in the initial (pre-1980s) steady-state.

The choice to commit crime involves weighing the costs and benefits of doing so. The benefits are common to all individuals in the economy: an instantaneous reward drawn from a common distribution. The costs are starkly different across individuals. While all face the same prison risk, what they lose by going to prison depends on their current state. In addition to forgone earnings in prison, they are also subject to human capital depreciation from non-employment and unemployment after leaving prison, both of which lower their future expected earnings. These opportunity costs are higher for individuals with high human capital or currently employed. Figure 4(a) shows the probability of committing crime conditional on receiving an opportunity. This probability decreases as human capital increases, and notably does so at a faster rate for the lower half of the human capital range. It is also slightly lower for employed people than for unemployed people.

Criminal capital, the prison flag, and age also contribute to criminality in the model. Figure 4(b) shows that individuals with high criminal capital are more likely to commit crimes and their criminality is less responsive to their human capital. This is partially

Moment	Data	Model
Incarceration - young and middle	0.59%	0.59%
Incarceration - old	0.09%	0.09%
Unemployment duration	20 weeks	20 weeks
Employment rate - young and middle	76.2%	76.5%
Recidivism rate (1 year)	19.9%	20.1%
Wage Ratio (criminals vs non)	86.4%	86.4%
criminal with prior	64.2%	59.3%
Regression coefficient- β^M	0.13	0.13
Regression coefficient- β^O	0.21	0.21
Regression coefficient- β^N	-0.005	-0.005
income persistency	0.96	0.96
income std	0.20	0.20

Notes: The Table shows a comparison of empirical and simulated moments. See Appendix for a detailed discussion for data sources on the empirical moments.

Table 3: Model Match

mechanical: high criminal capital types have more crime opportunities. Additional crime opportunities affect crime choices by lowering the value of human capital and employment, since high criminal capital individuals know they are more likely to commit crime and return to prison. This amplifies criminal behavior by further lowering the opportunity cost of crime. The prison flag also increases crime by lowering the likelihood an unemployed individual with a criminal record may find a job. However, we find this channel is quantitatively small.

	Criminals	Overall
Employment rate	74.8%	76.5%
Human capital	1.16	1.42
Prison Flag	59.3%	3.7%
Young and middle population	74.9%	34.0%

Table 4: Characteristics of Criminals

Table 4 shows how individuals who commit crimes in a given period in the stationary equilibrium differ from the overall population, comparing those with a prison flag to those without. Criminals are less likely to be employed, with lower human capital and younger, and more likely to have criminal record in their history. In the initial steady-state only 74.8% of criminals are employed, compared with 76.5% of the general population. The fact that most crime is committed by employed individuals may be surprising, but the same is true in the data.⁴¹ Part of the reason why is a difference in pay. The human capital (and wages) of criminals is on average 20% lower than that of the overall population (1.16 vs.

⁴¹The odds ratio of crime for employed individuals relative to unemployed individuals is 0.82 in the model. This is non-targeted and is actually a bit higher in the data at 0.86.

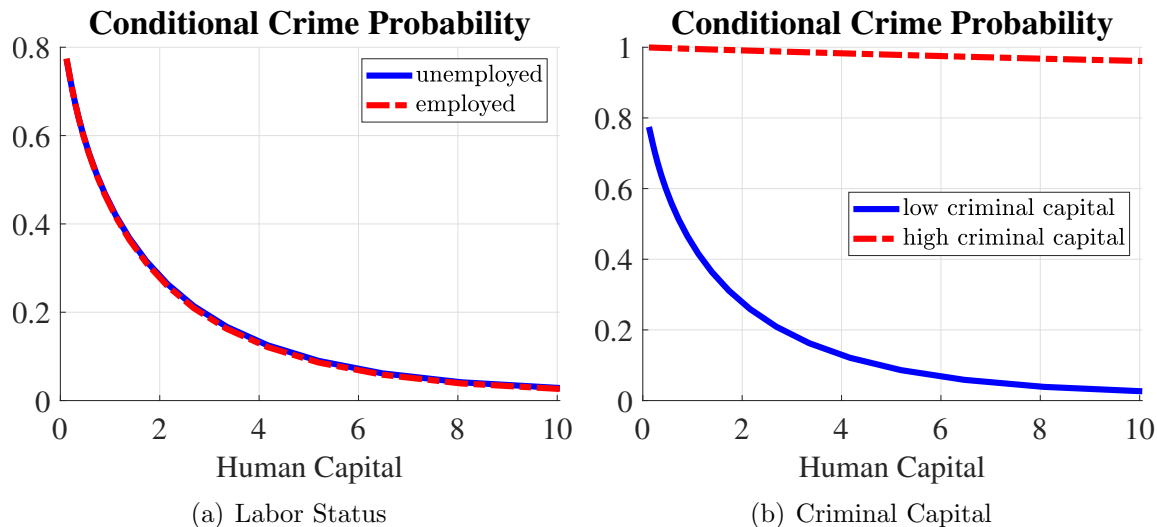


Figure 4: **Determinants of Crime:** The figure shows model generated crime probability conditional on receiving an opportunity as a function of human capital, labor market status and criminal capital for a middle-age agent.

1.42). As is well known in criminology, age is an important factor. More than three-quarters of the criminals are young and middle-aged individuals, whereas young and middle-aged individuals make up only one-third of the overall population.

Crime is highly concentrated, and the majority of the crimes are committed by a few individuals. The share of criminals with prior criminal history is almost 59.3%, although the share of this group in the overall population is only 3.7%. Only 1.6% of individuals commit 95% of all crimes in the initial steady-state. Given these stark statistics, the entry margin into crime and the persistence of crime for individuals who do enter will be key in the policy analysis of the dynamics of deterrence in the following sections.

A simple linear regression of individuals' probabilities of committing crime within the week is another statistic summarizing factors correlated with criminality (Table 5).⁴² The cells list the percentage-point change in crime probability associated with each independent variable. The average weekly crime probability in the economy is 0.1%, with a standard deviation of 0.5 percentage-points. The regression shows high wage is the main deterrent factor for crime. Employment has a deterrent effect, but it is statistically insignificant. The prison flag has a strong positive relationship with crime. This result further emphasizes how crime is concentrated in a few persistent individuals.

⁴²Specifically, the dependent variable is $\log(\frac{1}{p_c} + 1)$, where p_c is the probability of committing a crime conditional on receiving an opportunity. The independent variables are dummies for middle and old age, prison flag and employment, and log of wages.

	Estimate	SE	tStat	pValue
Age 25-34	0.07	0.01	10.68	0.0
Age 35-50	0.23	0.01	39.1	0.0
Prison Flag	0.08	0.01	7.81	0.0
Employed	-0.003	0.003	-0.68	0.49
ln(wage)	-0.12	0.0	-48.28	0.0
Constant	0.30	0.01	46.64	0.0

Table 5: Crime Elasticities

6 The Dynamics of Punitive Incarceration Reform

In this section, we study the effects of an increase in incarceration probability after committing a crime on aggregates like crime rates, incarceration rates, labor market variables, and inequality after the 1980s. Our intention is to mimic changes in punitive justice policy thought to be a key driver of the prison boom.⁴³ These policies, however, did not occur in isolation. Other factors shaping criminality evolved during this time as well. The first is the real wage stagnation of low-skilled workers that occurred through the late 1990s. The second is potential changes in crime rewards. While crime rewards cannot be directly observed, there is evidence that the spread of cocaine and associated gangs raised criminal involvement from the late 1980s through the mid 1990s (Blumstein (1995)). Our second motive is that these theories correct the deficiencies of and complement the strengths of a theory of unilateral change in incarceration policy. They magnify the impact on the incarceration rate and labor markets where incarceration policy alone quantitatively under-predicts trends from 1990 onward, and they counteract the decline in crime that is over-predicted by policy changes alone. It is necessary to consider all three changes together, as they will interact through the various channels in our model.

We calibrate the increase in the incarceration probability, the increase in the mean of the crime reward distribution, and the decrease in the productivity to match time trends in each: the incarceration rate, crime rate, and the employment rate over 1980-2010.⁴⁴ We feed the changes in these parameters linearly over the 25 years, and in each period the change is introduced as a surprise and permanent change. The calibration yields a seven-fold increase in the arrest probability, a 20% increase in the mean of the crime reward distribution, and a 30% decline in the average productivity of workers.

⁴³See Neal and Rick (2014), Blumstein and Beck (1999), Pfaff (2012), and Raphael and Stoll (2009), among others, for evidence that a change in admission conditional on crime was the main policy change. This probability changed little during crime waves, except for violent crime in the 1990s and in episodic instances of prison crowding in specific states. Median prison durations were relatively consistent.

⁴⁴More specifically, we fit a third-order polynomial for each data series, and targeted these smoothed data series.

6.1 Comparison of Initial and Final Steady States:

To see how a change in incarceration probability, π , affects the incarceration rate, define the probability of incarceration for an individual with current state s : $p^c(s; \pi) = \pi\eta(1 - H(\kappa^*(s)))$. The overall crime rate is $\int p^c(s; \pi)d\mu(s; \pi)$, where μ is the distribution of individuals across states and κ^* is the reservation crime reward. Increasing π affects the overall crime rate through three channels. The first is an arithmetic effect: the incarceration rate is the product of overall crime and π . The second is deterrence: higher π increases each individual’s choice of a reservation reward κ^* regardless of their state s . The final effect is how both π and all the endogenous responses in the model change the distribution of individuals across states μ . This includes the endogenous job creation response of the firms.

Whether an increase in π will increase or decrease incarceration rates is a quantitative issue. If the deterrence created by an increase in π is relatively small, then the arithmetic effect of a higher π can dominate and cause an increase in the incarceration rate. This typically generates a “Laffer curve” type of hump-shaped relationship between π and the incarceration rate. Incarceration rates are zero when $\pi = 0$ (no criminals go to prison) and when $\pi = 1$ (nobody commits crime). What is unusual in our model is that it is unclear that crime rates should fall. In simple settings, crime falls because all individuals raise their threshold $\kappa^*(s)$ and commit less crime. In our model, a prison experience worsens an individual’s state and makes them more likely to commit crime. In this way it is possible that an increase in π leading to an increase in incarceration could also increase crime.

The estimated trends of both lower productivity and higher reward for crime along the transition work to increase crime and incarceration. Table 6 shows the comparison of the initial and final steady-states.

Steady-State Variables	SS1 $\pi = 0.5\%$	SS2 $\pi = 2.9\%$
Incarceration	0.59%	0.96%
Crime Rate	1.4%	0.3%
Employment rate	76.5%	74.8%
Recidivism rate-1 year	20.1%	76.3%
Criminals with prison flag	59.3%	86.9%
Frac w/ high criminal capital	1.9%	0.7%
With prison flag	3.7%	2.0%
Share committing 95% of crimes	1.6%	0.1%
Wage ratio	86.4%	72.2%

Notes: The table shows a comparison of two steady states, one with $\pi = 0.5\%$ and one with $\pi = 2.9\%$, productivity 20% lower and crime reward 30% higher.

Table 6: Steady-State Comparison

Crime decreases from a rate of 1.4% to 0.3% across the steady-states in the simulation, as shown in Figure 5(a). The increase in the incarceration probability offsets the fall in crime, and as a result, the incarceration rate increases from 0.59% to 0.96%. It is notable, however, that crime becomes more concentrated within a fewer individuals. The share of the population responsible for committing 95% of aggregate crimes decreases from 1.6% to 0.1%. The fraction of individuals with a prison flag decreases from 3.7% to 2%, and the fraction of crime committed by individuals with prior conviction increases from 59.3% to 86.9%. These repeat offenders have a substantially higher recidivism rate even though their reservation reward increases, dominated by the large increase in incarceration probability. The wage ratio of criminals to the overall population decreases from 86.4% to 72.2%, which reflects the incentives to be more picky about crime.

The changes across steady states are due to changes in the policy functions of individuals and changes in the distribution of individuals. Figure 5(a) shows the changes in agents across states or types. An example of how policy functions change is shown in Figure 5(b). Observe that crime policy falls for all.

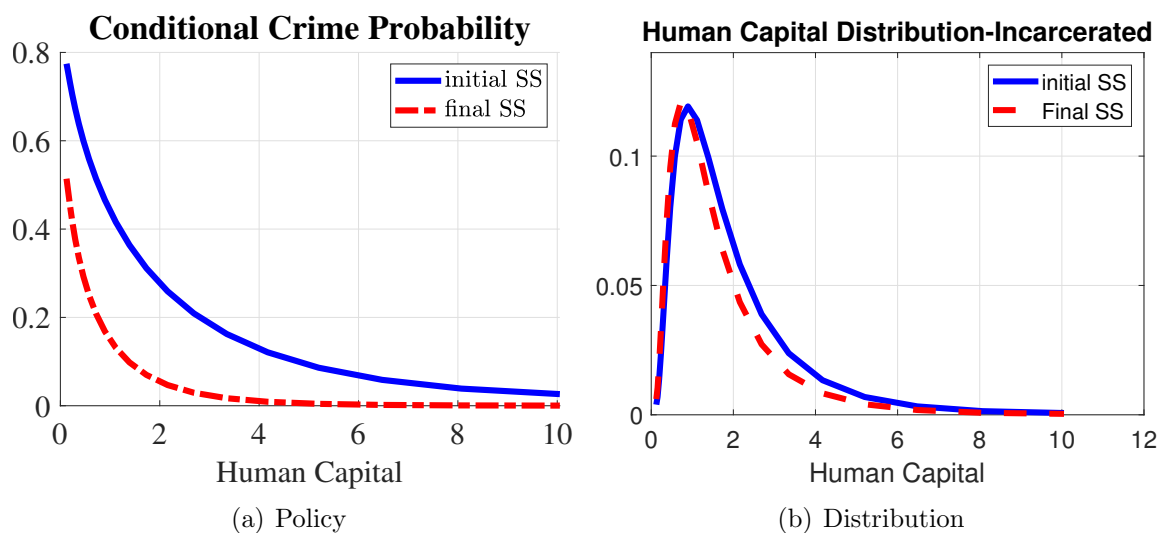


Figure 5: **Steady-State Comparison:** The left panel shows model-generated crime probabilities conditional on receiving an opportunity as a function of human capital for a middle-aged employed individual with low criminal capital and no prison flag across the initial and the final steady states. The right panel plots the distribution of human capital among the incarcerated across the initial and the final steady-states.

6.2 Transitional Dynamics:

The transition from the initial to final steady state can take several decades, and substantial policy costs occur along this transition. Figure 6 plots the transitional dynamics for incar-

ceration rate, crime rate and employment rate. It is not surprising that we match the overall pattern for each variable since we target them using the change in the arrest probability, productivity and crime reward.

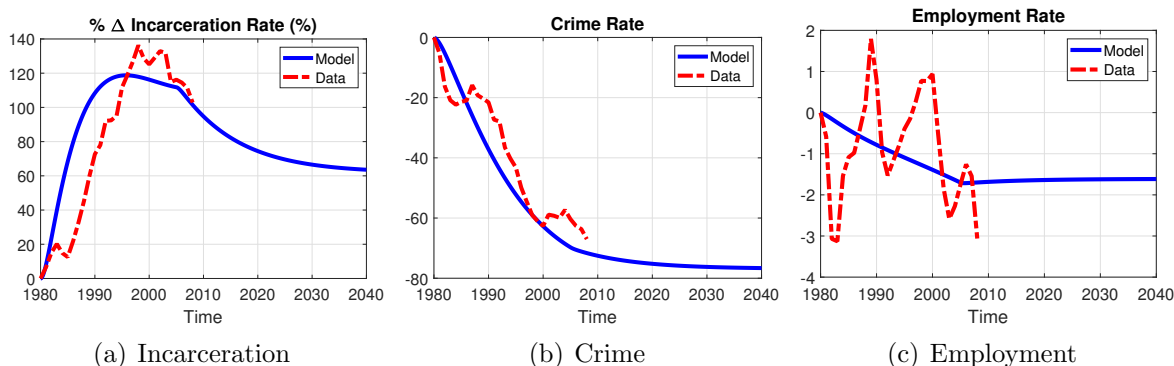


Figure 6: **Transitional Dynamics - Model vs. Data:** The figure shows the evolution of the incarceration rate, crime rate and employment rate along the transition. The left panel plots the total incarceration rate. The middle one plots the total crime rate, and the right panel plots the employment rate relative to their initial steady-state levels. The solid lines correspond to their model counterparts, whereas dashed lines correspond to the data.

Figure 6(a) shows the evolution of total incarceration rate along the transition, relative to the initial steady-state. It starts at 0.59%, more than doubles in 15 years, and then gradually declines to the new steady-state level of 0.96%.⁴⁵ This non-monotonic change in the incarceration rate happens despite the monotonic decline in the crime rate as captured in Figure 6(b). A similar relationship appears for total property and violent crime and overall incarceration rates in the U.S. during this time. A naive analysis may conclude that if crime is falling as incarceration rates fall, then punitive incarceration is not driving the fall in crime. Our theory shows the fault in this logic. Past moves towards more punitive incarceration policy follow dynamics where the full deterrent effect is delayed, and thus crime and incarceration decrease at a tipping point where the added dynamic deterrence overtakes the arithmetic increase in π .

Shapley-Owen Decomposition of How Shocks Shape the Trend. Figure 7 plots a Shapley-Owen decomposition of the trend into the three series we feed in: incarceration policy π , labor market productivity, and crime rewards. The contribution of a shock is computed by first calculating the contribution of that shock to the change in the variable of interest, compared with the initial steady-state in the presence and absence of the other

⁴⁵The acceleration in the incarceration rate after 2000 is largely due to the change in the denominator—the number of men without a high school degree is falling throughout this time but the education variable in the NCRP data is unreliable and we are unable to isolate admissions of that group.

two shocks. We take all possible permutations of them and then compute the weighted sum of each contribution of the shock in all the permutations according to the Shapley-Owen combinatorial formula (Shapley et al. (1953) and Owen (2014)).

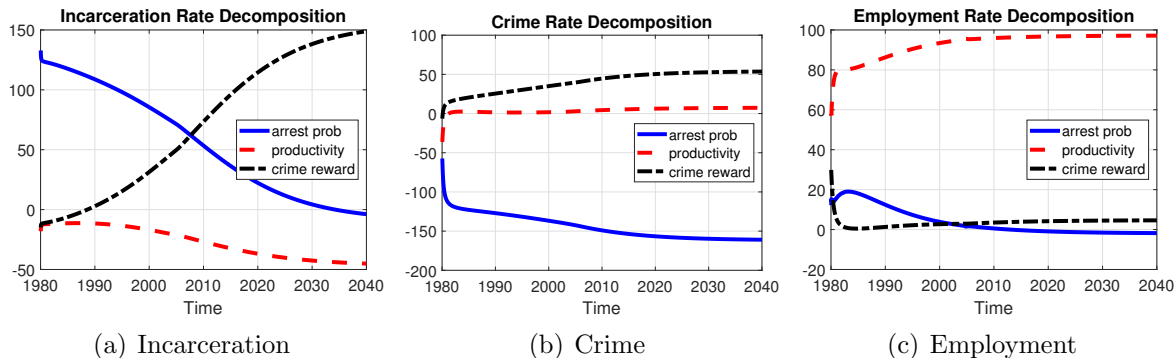


Figure 7: **Transitional Dynamics - Shapley-Owen Decomposition:** Solid lines show the contribution of the change in incarceration probability, dashed line shows the contribution of the change in the productivity, and finally the long-dashed line shows the contribution of the change in the crime reward. The left panel is for the incarceration rate, the middle panel is for the crime rate and the right panel is for the employment rate.

Figure 7(b) shows that the main driver of the evolution of the crime rate is the deterrence provided by an increase in the probability of incarceration for a crime. As policy becomes more punitive, individuals' crime rates decrease. In the absence of the other two shocks, we would expect the crime rate to drop another 60% in the long-run, but both productivity and inferred crime rewards work against changes in punitive policy and increase the crime rate. Changes in incarceration policy increase the incarceration rate in earlier periods but decrease it later on. This has to do with whether the decrease in crime is arithmetically large enough to offset the increase in π . Quantitatively, we find that in the absence of the other two shocks, the incarceration rate in the new steady state would be lower than in the initial steady-state. Finally, productivity is the main contributor to the change in the employment rate.

6.3 Dynamics of Deterrence:

This section seeks to understand the dynamic impacts of punitive policy by studying outcomes following changes in the probability of imprisonment for a crime in absence of the other factors we studied along the transition.

Incapacitation versus Deterrence. The criminology literature frames punitive policy as reducing crime through two channels: deterrence and incapacitation. Incapacitation lowers

crime by putting likely criminals in prison, where they cannot commit crime. Deterrence is when more punitive policies lower crime by deterring individuals from committing crime in the first place.

Two experiments in the structural model provide novel insights on how these effects unfold dynamically after a policy change. The incapacitation effect is isolated in the first by setting the time spent in prison to 0.⁴⁶ The deterrence effect is isolated in the second by fixing the decision rules of the individuals and firms at the initial steady-state level along the transition.

We start by looking at the effects on incarceration, shown in 8(a) . When incapacitation is eliminated, incarceration rates obviously go to zero. What can be seen in the difference between the benchmark (blue) and the line with incapacitation but without deterrence (black) is the impact of deterrence. Without deterrence, the incarceration rate follows the increase in π , reaching 2.5 times higher than the peak with deterrence and staying at that level forever.

The effects on crime, shown in 8(b), are more nuanced. The pink line has no incapacitation or deterrence but includes the higher impact of prison by giving those who are caught all the expected cumulative impacts of prison (prison flag and lower human capital) and sends them straight to unemployment. The fact that this line increases crime from the initial steady state shows the criminogenic effect of prisons. Lower human capital and losing employment causes ex-felons to choose more crime. The next line, the red dashed line, adds in deterrence but not incapacitation. Crime rises initially before falling later on. This emphasizes that the policy deters mostly through crime entry and it takes many years for the full deterrent impact on crime entry to be realized. It also shows how important incapacitation is, especially in the short run. The black-line shows incapacitation only. It always works to decrease crime and hits immediately with changes in π (recall π is fed in linearly).

The broad lesson is that the incapacitation effect is what reduces crime immediately after a change to more punitive policy but deterrence is what provides the majority of the decline in the long run, around 75% of the decline in our calibration. This finding suggests that changes in punitive justice policy may be improved by following a thought-out dynamic path. In particular, larger crime reduction could be achieved more quickly by lengthening prison duration in the short run, when incapacitation effects are key, and reducing them in the long run, when full deterrence has kicked in.⁴⁷

⁴⁶All of the expected cumulative effects of prison on human capital, the prison flag, and criminal capital from the baseline model are maintained.

⁴⁷Criminology studies have repeatedly found that lengthening prison sentences past a year or two provides virtually no additional deterrence. A recent example is Rose (2021).

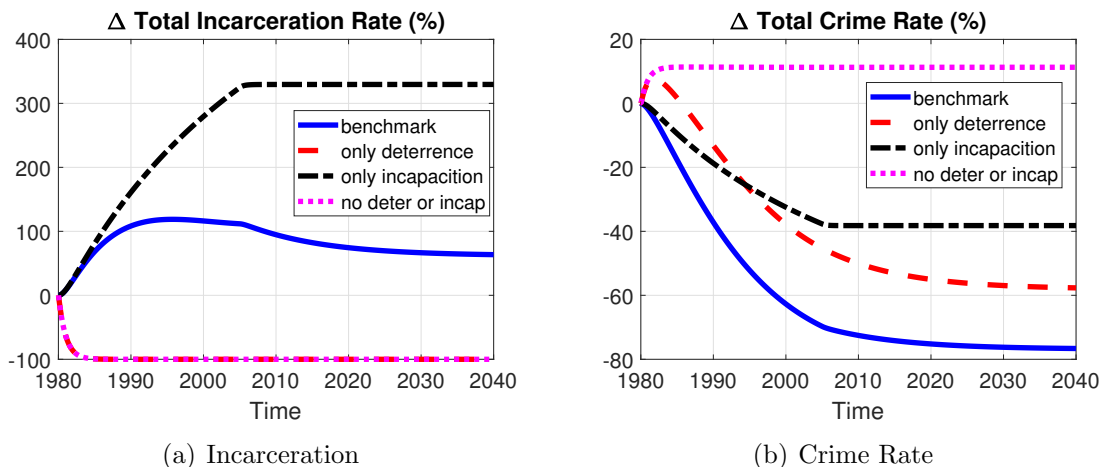


Figure 8: **Incapacitation vs Deterrence:** The figures compare the evolution of incarceration and crime rate along the transition without incapacitation or deterrence effects. The solid line is the benchmark economy. The red dashed line is the economy when incapacitation is eliminated. The black dashed line is the economy when deterrence is eliminated and all policy functions are fixed at the initial steady state. The pink dotted line eliminates both the incapacitation effect and the deterrence effect.

The Intensive and Extensive Margins of Crime: The decomposition of the transition highlighted that the crime entry decision is a key margin through which punitive incarceration policy provides deterrence. Another way to explore this theme is by investigating how the intensive and extensive margins of crime evolve. In other words, does crime become concentrated in fewer individuals, and do those fewer individuals commit more or less crime than earlier generations?

One measure of the extensive margin is how concentrated crime is across individuals. Figure 9(a) plots the evolution of the share of individuals responsible for a given fraction (80%, 90%, or 95%) of crimes along the transition. Crime unambiguously becomes more concentrated in fewer individuals. For example, the solid line shows a little over 1.6% of the population was responsible for 95% of crime at time zero and this falls to 0.1% of the population at the new steady state.

One measure of the intensive margin of crime is recidivism relative to incarceration probability π (9(b)). While it is true that recidivism increases over time (x4), it increases by less than would be arithmetically implied by the increase in π (x7). This means that the intensive margin is actually falling.

Combining these facts, we conclude that both the extensive *and* intensive margins are working together to provide the decrease in crime over time. Crime becomes more concentrated in fewer individuals *who actually do less crime each*. The increase in π always provides intensive deterrence through individuals raising their crime reward threshold for

any given state. The fact that the overall intensive margin (measured as recidivism) falls in the new steady state implies that the change in the distribution of individuals does not move the most criminally active to states that are bad enough to undo the deterrence provided by changes in policy functions. This is meaningful for practical policy because it implies that additional crime reduction of putting an additional person in prison (a pure marginal incapacitation effect) actually falls along the transition. This is a quantitative statement and could have plausibly gone the other way.

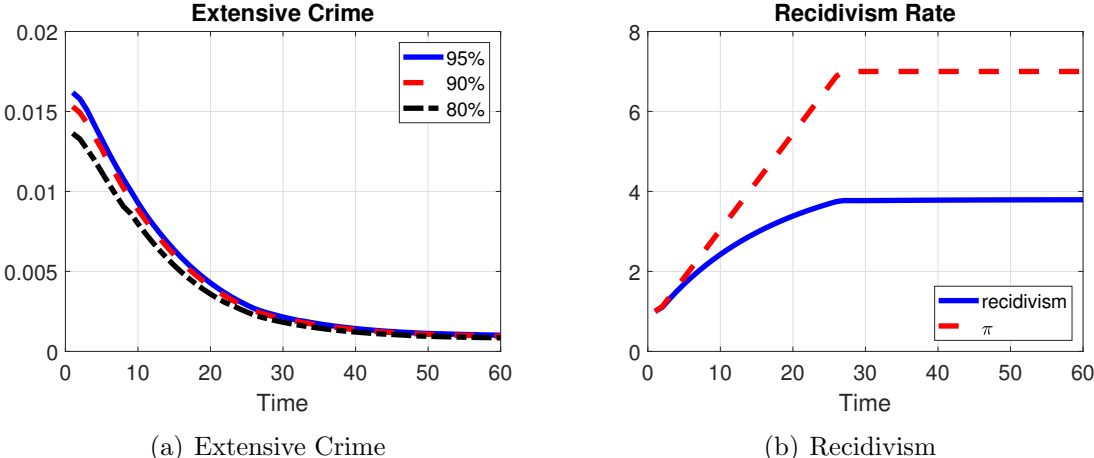


Figure 9: **Extensive Crime and Recidivism:** The left plots the measure of individuals committing certain shares of aggregate crime along the transition. The solid line is for 95% of crimes, the dashed line is for 90% of crimes, and the long-dashed line is for 80% of crimes. The right panel plots the one year recidivism rate, together with the arrest probability along the transition. Both the recidivism rate and arrest probability are normalized to their initial steady-state level.

Empirical measures are consistent with the model predictions that both the intensive and extensive margins of crime declined from 1980 to 2000. Table 7 shows the three year re-imprisonment rate has increased over time, but not as much as would be predicted by the seven-fold increase in observed prison admissions per crime. The same is true for the extensive margin: estimates of the percentage of people who would go to prison if they lived their lives entirely in a world of 2000-03 policy is higher than those living forever in 1974-79, but by less than would be implied by a seven-fold increase in π .⁴⁸ In these ways both crime entry (extensive margin) and repeat crime (intensive margin) offset the arithmetic impact of the increase in π on prison rates in the data.

⁴⁸The data show a 2.8-fold increase.

Total 3-Year Re-imprisonment			
Age	1983	1994	2000-2003*
18-24	64.0	41.0	48.8
25-34	32.6	40.3	49.6
35-64	27.0	35.6	44.3
Total (18-64)	30.7	39.3	47.7
Expected % of Population Incarcerated by age 35			
Year of Birth			
	1974-1979	1994	2000-2003*
	1.7	4.0	4.7

Table 7: Upper panel: 3-year Re-imprisonment Rate on a New Felony Charge, 1983 & 1994 Recidivism of Prisoners Released Series (United States Department of Justice. Office of Justice Programs. (2014)); *2000-2003: Florida only, (Bhati (2010)). Lower panel: estimated from Bonczar (2003) and authors' calculations in NCRP.

6.4 Cohort Effects.

Cohort effects are another source of insight into how crime entry decisions and criminal persistence of those previously involved in crime each drive the dynamics of deterrence. Figure 10(a) shows that the evolution of incarceration rate is different for different age groups in ways that are consistent with the empirical cohort evidence documented in Section 2. We show cohort effects in the model in Figure 10(d) and in the data in 2 The cohort effects in each are of remarkably similar shape and magnitude. Figure 10(d) shows the importance of cohort effects. The left line is the full empirical projection estimated from model data, and the right line is what is left after removing estimated cohort effects.

Figure 10(a) shows the incarceration rate of young individuals rises the least and approaches the new steady state the quickest. The incarceration rate of the oldest individuals rises the most and approaches the new steady state the slowest. Middle-aged individuals are somewhere in between. Our theory accounts for these facts with the importance of the crime entry decision combined with criminal persistence. The impact of new cohorts choosing not to enter into crime drives the dynamics of deterrence and enters into the incarceration rate of young individuals first. Over time, as new cohorts age, this impact hits the incarceration rates of middle-aged and then older age groups.

Stepping back, we note that these figures emphasize that the collateral costs of the change in policy are borne unequally across cohorts. Collateral costs of the prison system include costs released inmates face upon re-entry as well as the costs their families and communities face during and following their imprisonment.

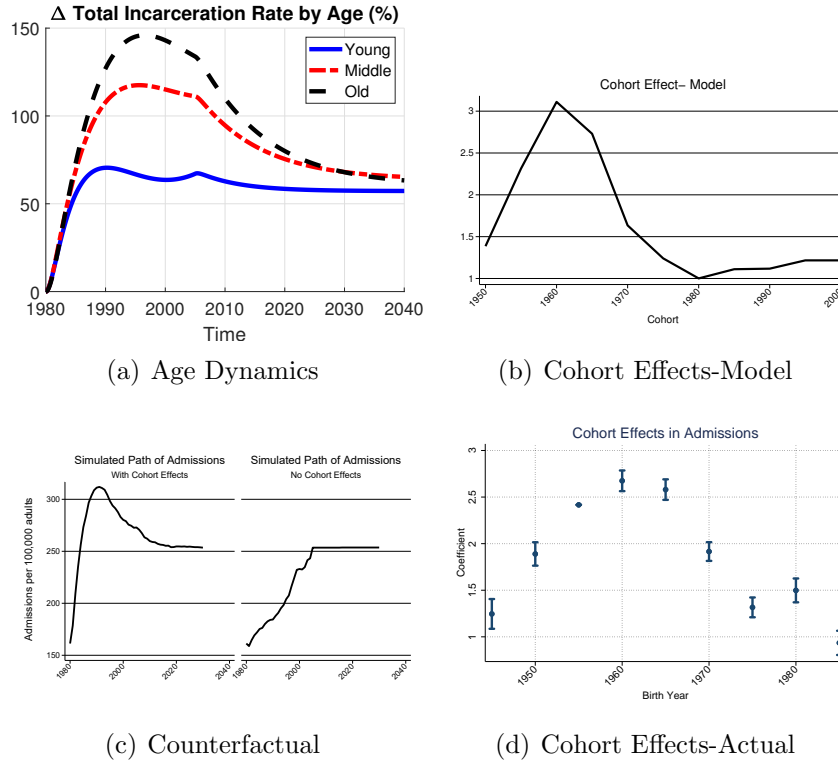


Figure 10: **Age Dynamics and Cohort Effects:** Top left: incarceration rate across different age groups as deviations from the initial steady-state. Right: estimated cohort effects from model (top) and data (bottom). Bottom left: empirical projection of model data with and without estimated cohort effects.

6.5 Incarceration Policy, Labor Markets, and Inequality

Punitive justice policy has frequently been cited as a potential contributor to the stalling in the closure of Black-White employment and income gaps beginning in the 1970s. While our model does not have race, it makes two clear predictions: moving towards more punitive policy has little impact on aggregate employment but increases inequality in employment and income.

Increased punitive incarceration policy alone had modest impact on aggregate employment after 1980. Figure 11(a) shows policy changes in isolation cause employment for people without prior incarceration, the majority of the population, to fall by 1.5 percentage points. The effect on individuals with a prior incarceration experience, however, is larger. Their employment rate falls by 4.0 percentage points, more than twice the drop for those who had not been to prison. Income dynamics, shown Figure 11(b), show similar disparate effects. While the income of the overall population decreases by 20%, individuals with a prior criminal record experience around 25% drop in their income.

Figure 11(b) is not the same as the causal impact of incarceration on earnings. To compute the causal impact in the model, we run a simulation where π , the market tightness,

and the initial distribution of individuals are fixed at the final steady state equilibrium but nobody actually goes to prison. We compare the same individuals in this world to themselves in the baseline and find the ones going to prison in the baseline earn 16%, 6.5%, 3.2%, 2.4% and 2.3%, respectively, in the first 5 years after they are released from prison than compared with their earnings paths in the counterfactual world. This is a little less than the causal estimate of prison on earnings of 13% less over five years found in Garin et al. (2024), but adds external validity to our quantitative results.

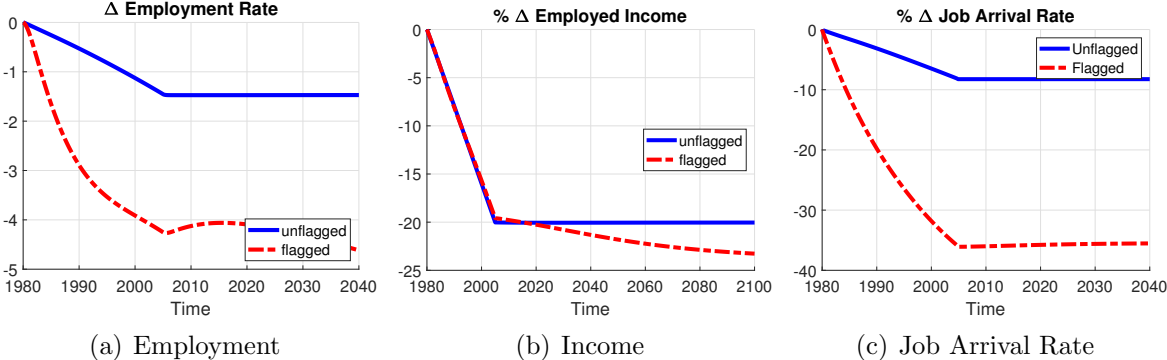


Figure 11: Employment, Income and Job Arrival Rates across Different Groups: The figures show the evolution of employment rate and income for individuals with and without a prison flag. The left panel is for employment, the middle panel is for income dynamics, and the right panel is for the job arrival rate of the middle-aged individuals. All are changes in percentage points relative to the initial steady-state level.

Employment is an equilibrium product of both individuals’ and firms’ responses to punitive policy. The response of individuals is simple. Every individual increases their threshold minimum reward required to commit a crime when the probability of incarceration for a crime rises. Quantitatively, changes in the policy function of individuals decrease overall criminal propensity by 75% in the new steady state. The remainder is a product of the changing distribution of agents across states and changes in the aggregate state, labor market tightness. Figure 5(a) shows employed individuals reduce their crime probabilities by around 75% in response to the policy change.⁴⁹

The only choice of a firm is whether to create a vacancy. This choice responds directly to punitive policy. All else equal, an increase in the probability of prison for a crime reduces the expected duration of a match with a worker and lowers a firm’s value of creating a vacancy. This choice also responds to policy indirectly through how it changes individual’s crime choices and the distribution of individuals. We have shown that both of these factors improve on average. Everybody chooses less crime, and the distribution averages higher

⁴⁹The magnitude of the drop in crime propensity is about the same for the unemployed.

human capital and lower criminal capital. All of these factors improve a firm’s value of creating a vacancy. We find that the direct impact of stricter policy dominates the indirect impacts of individuals’ response and distribution resulting in lower market tightness for both job seekers with a prison record and those without. The transition for the job arrival rate for the middle-aged individuals is shown in Figure 11(c). Individuals with a prison record flag have a larger decrease in market tightness because their criminality is higher than that of individuals without a prison record flag. An increase in π has more bite on their expected match duration, which falls by more than the expected duration for workers without a flag does. All of this occurs despite per capital crime rates falling across both markets.

Changes in market tightness can also feed back into criminal behavior and incarceration rates. We run three counterfactuals to quantitatively decompose how much of the firms’ response is due to the policy directly and how much is due to changes in individuals’ behavior, and to measure how much each piece affects crime, incarceration, and labor markets. Figures 12(a)-12(c) breaks down the benchmark transition into a counterfactual transition where only the policy function of individuals changes, a counterfactual transition where only the vacancy creation of firms changes, and a transition where neither change occurs (just a pure arithmetic impact of the policy).

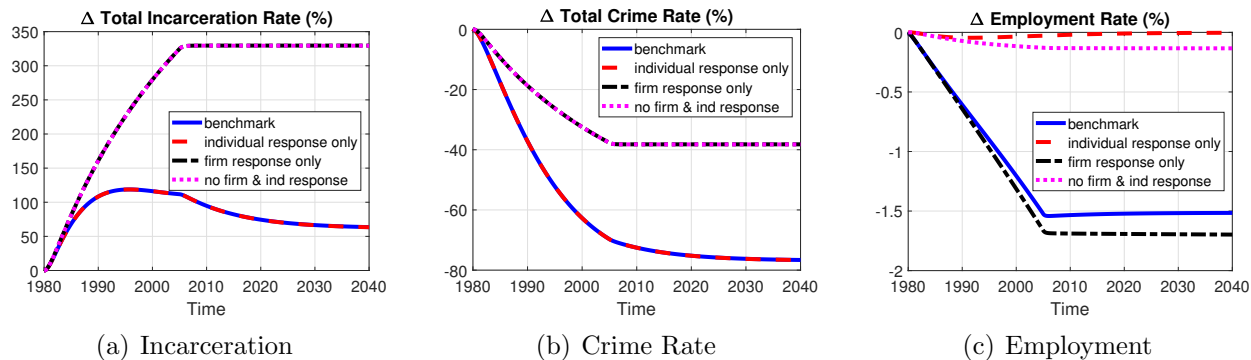


Figure 12: **Transitional Dynamics: Policy Decomposition:** The figures show the decomposition of incarceration, crime, and employment along the transition. The solid line is the benchmark economy. The dashed line is the economy when firms keep the same job creation level. The long dashed line is the economy when individuals keep their criminal policy as in the first steady-state. Lastly, the dotted line is the economy when firms keep the same job creation level, and individuals keep their crime choices as in the first steady-state.

In the last of these scenarios, incarceration rises by the same amount as the increase in π . Why then does crime fall when we omit the deterrence provided by changes in individual policy functions? This is due to the incapacitation effect of higher prison rates. Putting more of the most criminally active people in prison mechanically reduces crime. Comparing the pink dotted line to the benchmark shows that around half of the decrease in crime is from

incapacitation and half from deterrence (adding in the response of individuals' crime policy). Higher π also mechanically reduces the employment rate both through incapacitation and through higher churn of workers through prison to unemployment; it takes them time to find a new job.

The response of firms alone, shown in the black dashed line, has virtually no impact on crime and incarceration. This is not because firms don't respond; we have already shown vacancies and market tightness fall across the board. It is because crime is highly concentrated in individuals whose criminality responds little to changes in market tightness or employment status.

Panel (c) of Figure 13 shows that the firm response dominates in providing outcomes in labor markets. Without the firm response, the employment rate is reduced only by the higher share of the population in prison, but this is just 0.6% of the total population.

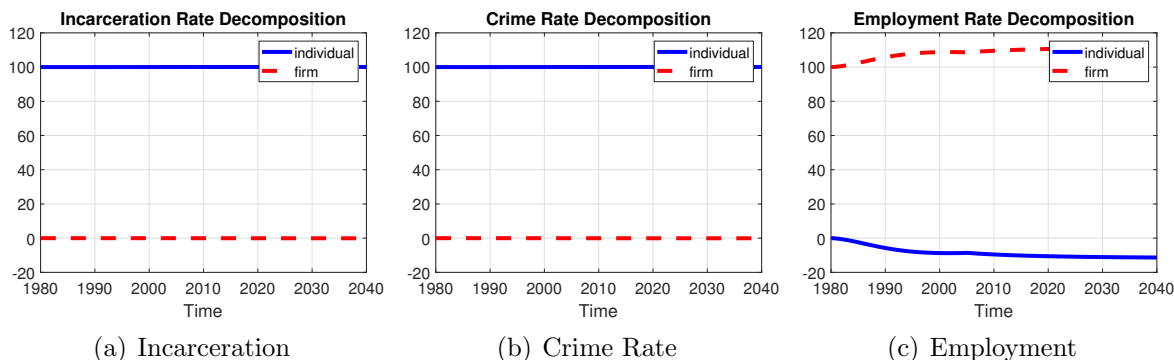


Figure 13: **Transitional Dynamics: Policy Decomposition:** The figures show the Shapley-Owen decomposition of individual and firm policy functions. The solid line is the contribution of individual criminal policy, and the dashed line is the contribution of firm vacancy policy. The left panel is for the incarceration rate, the middle panel is for the crime rate, and the right panel is for the employment rate.

6.6 The Importance of the Initial Steady State.

So far, we have taken a deep dive into the mechanisms that determine how crime and incarceration unfold over a transition following policy changes of magnitudes that comport with the 20th century U.S. prison boom. Yet, our structural approach allows more general lessons to be learned about how changes in punitive policy are likely to unfold. The most important lesson is that the elasticity of crime and incarceration in the short and long run—that is, the response of crime and incarceration to a marginal change in the probability of incarceration—depends on the initial steady state from which the policy is tightened.

Figure 14 plots short and long run elasticities across a range of initial values of π . A short-run elasticity will be defined as the change in crime and incarceration implied by changes

in the policy functions of individuals and firms to a 1% change increase in π , holding fixed the initial distribution of individuals' states. A long-run elasticity will be defined as the complete change in crime and incarceration at a new steady state with a 1% higher π .

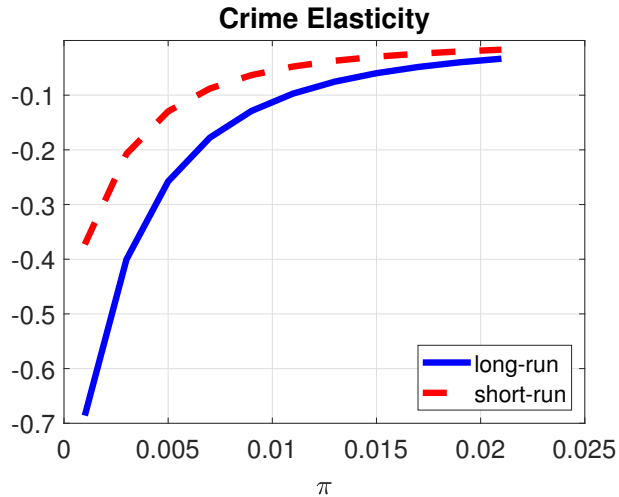


Figure 14: **Crime Elasticities:** The figure plots the aggregate crime elasticity in both the short-run and the long-run.

The long-run elasticities are always larger than the short-run elasticities, but this difference is larger when beginning at lower levels of π . Indeed, when starting from lax regimes, the long-run elasticity is twice the short-run. More punitive policy immediately makes all individuals choose less crime and critically deters the young from committing their first crime. Lower entry into the first crime is the key factor that lowers criminal capital in new cohorts and results in a distribution of types that commit lower crime. This is what distinguishes the slow-moving part of deterrence (long-run elasticity) from the instantaneous (short-run elasticity). This exercise shows that public policy evaluators should consider how a program or law separately affects crime entry and repeat offenders. Each effect carries independent information to evaluate the total long-run impact when only short-run information is available.

7 Alternatives to Criminal Capital

Criminal capital is a modelling tool that provides the persistence in criminal activity not accounted for by other features of the model, particularly labor market-related factors. It critically enables the model to match the intensive and extensive margins of crime in the population, and the hump-shaped cohort effects we identified along the transition to a more

punitive incarceration policy. In this section, we show that several alternative modelling assumptions are incapable of matching these features of the data as well as criminal capital can. We give each alternative the best shot possible by re-calibrating all parameters each time to best match initial targets. We also re-calibrate the shocks along the transition each time in attempts to match the evolution of the crime rate, incarceration rate and employment rate.⁵⁰

We remove criminal capital from the benchmark in the first experiment. Each subsequent experiment leaves criminal capital out and adds other features.

No Criminal Capital. Without criminal capital the model does not generate the concentration of crime in a few individuals with high recidivism, as in the data. The best fit one-year recidivism rate of this model is 0.5%, compared with its data counterpart of 19.9%. In the model only 7.7% of the criminals are among the repeated offenders, whereas the data counterpart is 64.2%. Simply put, crime is too widespread when considering pecuniary factors alone. Cohort effects are monotone along the policy transition.

Higher Human Capital Depreciation for High Criminal Capital Types. The model is slightly improved but still produces crime that is far too widespread and a recidivism rate that is far too low. The one-year recidivism rate is 0.5%, and the share of repeated offenders among criminals becomes 7.8%. Cohort effects are monotone along the policy transition.

Better or More Crime Opportunities for the High Criminal Capital Types. A version with a different mean of the distribution of crime opportunities for high criminal capital types also fails in matching the intensive and extensive margins of crime. The one-year recidivism rate is 0.75%, and the share of repeated offenders among criminals as 8.7%. An alternative model where high criminal capital types have a higher arrival rate of the same, not better, opportunities can better match the concentration of crime but implies a higher wage for criminal types than that for non criminal types which is the opposite of in the data (1.02 in the model vs. 0.86 in the data). Neither model generates the hump-shaped cohort effects similar to the data.

Ex-ante, Permanent Heterogeneity in Criminal Capital. Similar to the model with more crime opportunities for criminals, this model matches all the moments reasonably well except the wage ratio of criminals and non-criminals (0.99 in the model vs. 0.86 in the

⁵⁰We briefly explain these alternatives below and refer readers to the Online Appendix for a detailed explanation of each calibration.

data). Cohort effects are monotone along the policy transition. It also cannot generate the hump-shaped cohort effects similar to the data.

Higher Arrest Probability for the Incarcerated. This model is the next best after the benchmark in replicating the initial steady state targets but does not provide cohort effects in the transition.

This exercise revealed the key modelling features required to match the data targets. First, crime opportunities for recidivists can't be too rewarding, or else they would require a high wage for employment, contrary to the data. Second, criminal capital that is orthogonal from human capital is key to decoupling the otherwise strong relationship between labor market factors and criminality. This is important because the relationship between these factors is weak in the data. Finally, the punishment for crime, including labor market scarring through human capital depreciation, cannot be too costly. If the costs are too large, the model requires crimes to arrive infrequently but with a high reward that almost all agents would take. This effectively matches the crime rate in the data with a near exogenous shock, causing the model to miss the concentration of crime in fewer serial criminals.

8 A Comparison to Violent Crime

The dynamics of violent crime provide an example of the breadth of applications of this theoretical framework.⁵¹ We re-calibrate the model using targets from our criminal justice data limiting our sample to those who's primary offense was a violent one. Table 8 and Table 9 summarize notable differences between the violent and property crime calibrations. It also shows the model is capable of replicating targeted moments of violent crime.

Calibrated Parameter Values			
Parameter	Explanation	Value	
		Property	Violent
η^1	crime arrival rate	0.07%	0.05%
ζ^3	rehabilitation shock	0.39%	0.35%
ν	prob of being high criminal	0.23	0.15
$\eta_a^{1,hc}$	high criminal crime arrival rate	1.57	0.65
μ^k	mean crime reward	0.61	1.44

Table 8: Calibration to Violent Crime, with Comparison to property crime.

⁵¹Replications of all tables and figures located in this text for the case of violent crime are available in the Online Appendix.

Targeted Moments and Fit			
Moment	Data	Model	Property Crime
Incarceration - young and middle	0.44%	0.44%	0.59%
Incarceration - old	0.09%	0.09%	0.09%
Recidivism rate (1 year)	13.5%	13.4%	19.9%
Wage ratio	87.7%	87.7%	86.4%
Criminal with prior	53.7%	54.7%	64.2%

Table 9: Model fit, with comparison to property crime.

While the targeted moments may not appear to be very different, the model predicts violent crime is far more inelastic than property crime to changes in the incarceration probability π . The difference in elasticity between crime types is larger when starting in lax regimes (low levels of π). The crime entry decision is key to why violent crime has a lower elasticity to policy than property crime has. Violent crime has a recidivism rate that is two-thirds that of property crime, which leads to a higher share of crimes committed by first time offenders. Punishment is also higher for violent crime: the probability of incarceration is higher, and prison spells last longer. This leads the calibration to choose a crime reward process with an arrival rate that is two-thirds that of property crime but with a mean that is twice as large. At that reward level, the majority of individuals take a crime opportunity when it arrives, versus more discretion, for the lower rewards for property crime. In this sense, violent crimes look like “crimes of passion” that are relatively inelastic to individual characteristics and, subsequently, policy.

Empirical evidence supports the model’s contrasting predictions for the evolution of violent and property crime over 1980-2010. At the aggregate level, the model predicts the response of violent crime to policy is less dynamic than property crime: as shown in Figure 15, the change in the short run is similar to the new steady state. At the micro level, the model’s predicted cohort effects are monotone for violent crime, and non-monotone for property crime. This juxtaposition is consistent with the empirical evidence in Figure 2.

9 Conclusion

We argued that dynamics are critical when evaluating changes in punitive incarceration policy, because of criminal persistence. The majority of felonies in the United States involve individuals with prior criminal records, whose crime choices are less elastic to policy changes than the choices of those without records. The deterrent impact of more punitive incarceration materializes gradually and is strongest for crime entry margins pertaining to young and new generations. We presented novel empirical evidence on cohort effects consistent with

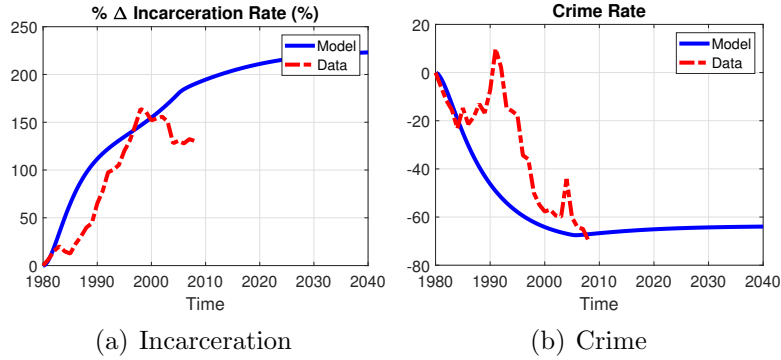


Figure 15: **Transitional Dynamics of Violent Crime- Model and Data**

this idea.

The dynamic model developed sheds further light on the sources of criminal persistence. It replicates salient features of criminal behavior that pecuniary motives alone could not explain: high recidivism rates, even among the employed and elderly, and cohort dynamics following the 1980s policy changes. We learned that unemployment and low human capital are instrumental in the choice to engage in crime, but criminal capital and, to a lesser extent, employment discrimination drive persistence after youth. Cumulatively, most crime is committed by a few individuals with lengthy criminal records, for whom pecuniary factors provide little deterrence.

In analyzing the impact of increased punitive incarceration akin to 1980s policy changes, we arrived at two substantive conclusions for property crime. First, the change in incarceration policy alone was a minor contributor to trends in low-skilled labor markets and aggregate incarceration from 1990 onward but a major contributor to crime reduction and increased inequality within low-skilled populations. Second, the transition after a policy change follows nuanced, multi-decade dynamics. Immediate incapacitation of the most active criminals drives initial incarceration increases. Subsequently, individuals cycling through prison re-enter the population with worsened labor market prospects and higher criminality. Full deterrent effects manifest only as new cohorts are born under the new policy, who choose lower crime and higher labor force attachment from youth. Applying the model to violent crime yields contrasting results, as it is less persistent and less elastic to policy changes, responding less but more immediately, with near-zero cohort effects.

While far from the final word on these important issues, we argue that dynamics should be addressed in future work. Interpretations of econometric inference should consider that short-run policy effects can differ dramatically from long-run effects, as demonstrated. Our structural model complements econometric inference by interpreting short-run effects to predict dynamic paths. Considering dynamics also introduces opportunities to improve

policies. When crime is more persistent, as with property crime, crime reduction immediately after a policy change comes almost entirely from incapacitation effects, while deterrence effects on crime entry build over time. These conclusions should encourage the study of dynamic punitive policies specifying paths for multiple levers: the probability and duration of incarceration, as well as differential penalties for new and repeat offenders.

References

- BECKER, G. S. (1968): “Crime and Punishment: An Economic Approach,” *Journal of Political Economy*, 76, 169–217.
- BHATI, A. S. (2010): “Criminal Recidivism in a Large Cohort of Offenders Released from Prison in Florida, 2004-2008,” .
- BLUMSTEIN, A. (1995): “Youth violence, guns, and the illicit-drug industry,” *J. Crim. L. & Criminology*, 86, 10.
- BLUMSTEIN, A. AND A. J. BECK (1999): “Population growth in US prisons, 1980 – 1996,” *Crime and justice*, 26 (January), 17–61.
- BLUMSTEIN, A. AND J. COHEN (1973): “A Theory of the Stability of Punishment,” *Journal of Criminal Law and Criminology*, 64 (2), 198 – 207.
- BONCZAR, T. P. (2003): “Prevalence of Imprisonment in the US Population, 1974-2001,” Special Report NCJ 197976, Bureau of Justics Statistics.
- BURDETT, K., R. LAGOS, AND R. WRIGHT (2003): “Crime, inequality, and unemployment,” *American Economic Review*, 93, 1764–1777.
- BUREAU OF JUSTICE STATISTICS (2011): “Recidivism Among Released Prisoners, 1983: [United States],” Ann Arbor, MI: Inter-university Consortium for Political and Social Research [distributor].
- BURNHAM, R. AND H. BURNHAM (1999): *United Nations World Surveys on Crime Trends and Criminal Justice Systems, 1970-1994: Restructured Five-Wave Data*, Inter-university Consortium for Political and Social Research.
- BUSHWAY, S. D. (2011): “So policy makers drive incarceration—Now what?” *Criminology & Public Policy*, 10, 327–333.
- CAUCUTT, E. M., N. GUNER, AND C. RAUH (2021): “Incarceration, Unemployment, and the Racial Marriage Divide,” .
- DOLEAC, J. L. AND B. HANSEN (2020): “The unintended consequences of ban the box: Statistical discrimination and employment outcomes when criminal histories are hidden,” *Journal of Labor Economics*, 38, 321–374.

- EISEN, L.-B. AND J. CULLEN (2016): “Update: Changes in State Imprisonment Rates,” Brennan Center for Justice.
- ENGELHARDT, B. (2010): “The effect of employment frictions on crime,” *Journal of Labor Economics*, 28, 677–718.
- ENGELHARDT, B. E., G. ROCHETEAU, AND P. RUPERT (2008): “Crime and the labor market: A search model with optimal contracts,” *Journal of Public Economics*, 92, 1876–1891.
- FELLA, G. AND G. GALLIPOLI (2014): “Education and crime over the life cycle,” *Review of Economic Studies*, 81, 1484–1517.
- FU, C. AND K. I. WOLPIN (2018): “Structural Estimation of a Becker-Ehrlich Equilibrium Model of Crime: Allocating Police Across Cities to Reduce Crime,” *Review of Economic Studies*, 85, 2097–2138.
- GARIN, A., D. K. KOUSTAS, C. MCPHERSON, S. NORRIS, M. PECENCO, E. K. ROSE, Y. SHEM-TOV, AND J. WEAVER (2024): “The impact of incarceration on employment, earnings, and tax filing,” Tech. rep., National Bureau of Economic Research.
- GOTTFREDSON, M. R. AND T. HIRSCHI (1990): *A general theory of crime.*, Palo Alto, CA: Stanford University Press.
- HINDELANG (2016): “Sourcebook of Criminal Justice Statistics,” Tech. rep.
- HUANG, C.-C., D. LAING, AND P. WANG (2004): “Crime and Poverty: A Search-Theoretic Approach,” *International Economic Review*, 45, 909–938.
- İMROHOROĞLU, A., A. MERLO, AND P. RUPERT (2004): “What Accounts for the Decline in Crime,” *International Economic Review*, 45, 707–729.
- LAGAKOS, D., B. MOLL, T. PORZIO, N. QIAN, AND T. SCHOELLMAN (2016): “Life-cycle wage growth across countries,” *Journal of Political Economy*, 126, 797–849.
- LOCHNER, L. (2004): “Education, Work, And Crime: A Human Capital Approach,” *International Economic Review*, 45, 811–843.
- MCCRARY, J. (2010): “Dynamic perspectives on crime,” in *Handbook on the Economics of Crime*, Edward Elgar Publishing.
- NEAL, D. AND A. RICK (2014): “The prison boom and the lack of black progress after Smith and Welch,” Tech. rep., National Bureau of Economic Research.
- OWEN, A. B. (2014): “Sobol’indices and Shapley value,” *SIAM/ASA Journal on Uncertainty Quantification*, 2, 245–251.
- PETTIT, B. (2012): *Invisible Men: Mass Incarceration and the Myth of Black Progress*, Russell Sage Foundation.

- PFAFF, J. F. (2012): “The micro and macro causes of prison growth,” *Georgia State University Law Review*, 28, 1239.
- PISSARIDES, C. A. (1985): “Short-Run Equilibrium Dynamics of Unemployment, Vacancies, and Real Wages,” *American Economic Review*, 75, 676–690.
- RAPHAEL, S. AND M. A. STOLL (2009): “Why are so many Americans in prison?” *Do prisons make us safer*, 27–72.
- ROGERSON, R., R. SHIMER, AND R. WRIGHT (2005): “Search-theoretic models of the labor market: A survey,” *Journal of economic literature*, 43, 959–988.
- ROSE, E. K. (2021): “Does banning the box help ex-offenders get jobs? Evaluating the effects of a prominent example,” *Journal of Labor Economics*, 39, 79–113.
- SCHULHOFER-WOHL, S. AND Y. C. YANG (2016): “Modeling the evolution of age and cohort effects,” *Dynamic Demographic Analysis*, 313–335.
- SEVIGNY, E. L. AND J. P. CAULKINS (2004): “Kingpins or mules: An analysis of drug offenders incarcerated in federal and state prisons,” *Criminology & Public Policy*, 3, 401–434.
- SHAPLEY, L. S. ET AL. (1953): “A value for n-person games,” Tech. rep.
- SHIMER, R. (2005): “The cyclical behavior of equilibrium unemployment and vacancies,” *American economic review*, 25–49.
- STEFFENSMEIER, D. J., E. A. ALLAN, M. D. HARER, AND C. STREIFEL (1989): “Age and the distribution of crime,” *American journal of Sociology*, 94, 803–831.
- STORESLETTEN, K., C. I. TELMER, AND A. YARON (2004): “Cyclical dynamics in idiosyncratic labor market risk,” *Journal of political Economy*, 112, 695–717.
- UNITED STATES DEPARTMENT OF JUSTICE. OFFICE OF JUSTICE PROGRAMS. (2014): “Recidivism of Prisoners Released in 1994.” .